

November, 1963

Risø Report No. 70

Interpretation of Low-Activity Counting

by

J. Thomas

The Danish Atomic Energy Commission
Research Establishment Risø
Electronics Department

Abstract

This report deals with the problem of estimating the unknown activity of a radioactive source. The estimate is based on the information that the radioactive source plus the background of the detector have given the count number n_1 in a certain time interval, and that the background alone has given the count number n_2 in a period of the same length. Further it is assumed that the activity of the radioactive source is positive or equal to zero.

The mean value and the standard deviation of the probability distribution of the unknown activity are derived from these assumptions, and tabulations are made and approximative expressions derived for practical use.

RISØ BIBLIOTEK

5100015300117



This report is a translation of the main part of an original Danish treatise which received a proxime accessit awarded by the University of Copenhagen in 1962. Publication has been made possible by the kind permission of the University.

Contents

	Page
Introduction	5
Theory of the Counting of Radioactive Decays	5
The Concept of Activity	9
Derivation of $H(n; a)$ from the Poisson Distribution	13
Repetition of Measurement of a Radioactive Source with Constant Activity	14
$H(n; a)$ is Asymptotically Normal	16
The Difference between Two Activities, a_1 and a_2	17
Background Counts	18
Mean Value and Standard Deviation in the Case with a Background ...	22
Approximative Expressions for $E[a]$ and $D[a]/E[a]$	25
Summary	29
Acknowledgments	31
Appendices	
1. Theorem on Semi-invariants	32
2. Calculation of Expressions for $E[a]$ and $D[a]$	33
3. Summation of $\sum_{r=0}^{n_1} \frac{(n_2+r)!}{r!} (\frac{1}{2})^r$	38
Table 1	42
References	60

Introduction

The intention of this report is to discuss the theory of interpretation of counting experiments. In such experiments the activity of the radioactive source will be the unknown parameter which has to be estimated on the basis of the observations.

In the treatment given by Bateman¹⁾ the activity was assumed to be known, and the variance (standard deviation) of the series of count numbers from a repeated experiment was calculated. However, the observed count number from a single experiment does not in fact exhibit any spread, but the spread is attached to our estimate of the activity which was the primary source of that count number.

If the problem is to estimate the count number in a subsequent experiment with the same radioactive source, the old treatment can be used, but only after the activity of the source has been estimated. The result of this problem is then found by an integration of all possible values of the still unknown activity.

Theory of the Counting of Radioactive Decays

In order to obtain a general description of the decays in a radioactive source, the decay of one atom in the source may first be considered. We assume that each atom has a constant decay probability, λ (equal for all atoms in the source). It is defined in such a way that the probability that the atom decays in a small time interval, dt , is proportional to the length of dt and independent of how long the atom has existed. The decay probability, λ , is then the proportionality factor. We here assume that at any time the atom either exists or has decayed, so that there is no definite time interval during which the atom is decaying.

From this it is possible to give the probability $D(\lambda; t)$ that an atom with the decay probability λ decays in the time t by

$$D(\lambda; t) = \int_0^t \{1 - D(\lambda; \tau)\} \lambda d\tau, \quad (1)$$

$1 - D(\lambda; \tau)$ being the probability that the atom does not decay in the interval $(0, \tau)$. This equation has the solution

$$D(\lambda; t) = 1 - e^{-\lambda t}, \quad (2)$$

and we see that

$$\lim_{t \rightarrow \infty} D(\lambda; t) = 1, \quad (3)$$

which means that sooner or later the atom will decay; however, there is a positive probability, $e^{-\lambda t}$, that the atom has not decayed in the course of a finite time interval, t , however large t may be.

If a source with N equal atoms - i.e. atoms with the same decay probability, λ - is given, the calculation of the probability, $G(N, \lambda, t; n)$, that n of the N atoms decay in the time t , is an elementary problem of mathematical statistics, especially if the decay of one atom is independent of the decay of the other atoms²⁾. Except for a few extraordinary cases, this assumption may be considered correct, and thus $G(N, \lambda, t; n)$ is given by a binomial distribution

$$\begin{aligned} G(N, \lambda, t; n) &= \frac{N!}{n!(N-n)!} (D(\lambda; t))^n (1-D(\lambda; t))^{N-n} \\ &= \frac{N!}{n!(N-n)!} (1-e^{-\lambda t})^n (e^{-\lambda t})^{N-n}, \end{aligned} \quad (4)$$

$D(\lambda; t)$ being the probability that the individual atom decays in the time t .

Normally, the detectors used for detecting radioactive decays are not 100 per cent. effective. If the efficiency of such a detector (equal to the probability that it detects the decay of a single atom) is denoted $\epsilon \leq 1$, the probability, p , of the single atom decaying in the time t and being detected is given by

$$p = \epsilon \cdot D(\lambda; t), \quad (5)$$

and the probability, $G(N, p; n_0)$, of $n_0 \leq n \leq N$ of the N atoms decaying and being detected in the time t is then given by

$$G(N, p; n_0) = \frac{N!}{n_0!(N-n_0)!} p^{n_0} (1-p)^{N-n_0}. \quad (6)$$

Now it is easy to calculate the mean value of the count numbers observed, which appears when we repeat the experiment with known N and p . But in practice this problem very seldom occurs, and in the case of determination of the "strength" of the source it is only N which is unknown, while p is given by the efficiency of the counter, the decay con-

stant for the atom and the observation time t , and n_0 is the number of counts observed.

If it were experimentally possible to set $p = 1$, N would be identical with n_0 , but this would require counting for an infinite period with a detector 100 per cent. effective.

Normally $p < 1$, and therefore all values of $N \geq n_0$ are possible, but not equally probable, and the task must be to determine the probability, $H(n_0, p; N)$, that the source contained exactly N atoms at the time $t = 0$ when n_0 decays have been observed in the time t . As p is considered a known constant, it will in the following be stated as an index instead of a variable, i. e.

$$H(n_0, p; N) = H_p(n_0; N)$$

and

$$G(N, p; n_0) = G_p(N; n_0).$$

The method of solving this kind of problem is given by, among others, M. Annis et al.³⁾, who introduce the a priori probability, here $P(N)$, that the number of atoms in the source is N independently of the number of observed counts (i. e. before any observation has been made), and the probability $Q(n_0)$ that the number of counts observed is n_0 independently of N .

The simultaneous probability, $S_p(N, n_0)$, that the number of atoms at the time $t = 0$ is N and that the number of observed decays in the time t is n_0 is now given in two ways:

$$\begin{aligned} S_p(N, n_0) &= Q(n_0) \cdot H_p(n_0; N) \\ &= P(N) \cdot G_p(N; n_0), \end{aligned} \tag{7}$$

from which we have

$$H_p(n_0; N) = \frac{P(N) \cdot G_p(N; n_0)}{Q(n_0)}. \tag{8}$$

If the two expressions for $S_p(N, n_0)$ are summed up over all possible values of N (i. e. $N \geq n_0$), we obtain

$$\sum_{N=n_0}^{\infty} Q(n_0) \cdot H_p(n_0; N) = Q(n_0)$$

$$= \sum_{N=n_0}^{\infty} P(N) \cdot G_p(N; n_0), \quad (9)$$

as, in order to be a probability, $H_p(n_0; N)$ must fulfil the condition

$$\sum_{N=n_0}^{\infty} H_p(n_0; N) = 1. \quad (10)$$

If the expression found for $Q(n_0)$ is inserted in formula (8), we have

$$H_p(n_0; N) = \frac{P(N) \cdot G_p(N; n_0)}{\sum_{N=n_0}^{\infty} P(N) \cdot G_p(N; n_0)}, \quad (11)$$

where $G_p(N; n_0)$ is known (eq. 6). In order to proceed it is necessary to assume an a priori probability $P(N)$. We know that $P(N) = 0$ for $N < 0$; but sources produced with any number of atoms might be possible, and consequently it is reasonable to assume any value of $N \geq 0$ equally probable.

Hence

$$H_p(n_0; N) = \frac{G_p(N; n_0)}{\sum_{N=n_0}^{\infty} G_p(N; n_0)} \quad (12)$$

$$= \frac{\binom{N}{n_0} p^{n_0} (1-p)^{N-n_0}}{\sum_{N=n_0}^{\infty} \binom{N}{n_0} p^{n_0} (1-p)^{N-n_0}}$$

$$= \binom{N}{n_0} p^{n_0+1} (1-p)^{N-n_0}. \quad (13)$$

The mean value, $E[N]$, of this distribution has the direct meaning that on the basis of the observed count number n_0 the probability distribution found for the true number of atoms at the time $t = 0$ has $E[N]$ as its centre of gravity. The mean value is found by

$$\begin{aligned}
 E[N] &= \sum_{N=n_0}^{\infty} N \cdot H_p(n_0; N) \\
 &= \sum_{N=n_0}^{\infty} N \binom{N}{n_0} p^{n_0+1} (1-p)^{N-n_0} \\
 &= \frac{n_0+1}{p} - 1.
 \end{aligned} \tag{14}$$

As a measure of the concentration in probability about this mean value we can use the standard deviation $D[N]$ (often written $\sigma(N)$), given by

$$\begin{aligned}
 D^2[N] &= E[N^2] - E^2[N] \\
 &= \frac{(n_0+2)(n_0+1)}{p^2} - 3 \frac{n_0+1}{p} + 1 - \left\{ \frac{(n_0+1)^2}{p^2} - 2 \frac{n_0+1}{p} + 1 \right\} \\
 &= \frac{n_0+1}{p^2} (1-p).
 \end{aligned} \tag{15}$$

We see that

$$\begin{aligned}
 E[N] &\rightarrow n_0 \\
 \text{and } D[N] &\rightarrow 0 \\
 \text{for } p &\rightarrow 1.
 \end{aligned}$$

The Concept of Activity

If the half-life of the radioactive isotope (equal to $\ln 2/\lambda$) is very long compared with the observation time, t , then the detection probability

$$p = \epsilon \cdot (1 - e^{-\lambda t}) \tag{5}$$

will be very small, and the number of atoms in the radioactive sample has to be very large to make observation of a decay in the time t possible. In order not to operate with such very large and very small quantities, it is

convenient to define the activity a of the source by

$$\begin{aligned} a &= N \cdot p \\ &= N \cdot \epsilon \cdot (1 - e^{-\lambda t}) \\ &\sim N \cdot \epsilon \cdot \lambda \cdot t \end{aligned} \tag{16}$$

for very small values of the product $\lambda \cdot t$.

a is still a discrete variable, and the probability, $H(n; a)$, of having a certain a when n decays have been observed in the time t is given by eq. (13), a synonymous relation between N and a being given by eq. (16). Thus we find

$$H(n; a = N \cdot p) = \binom{N}{n} \cdot \frac{a}{N}^{n+1} \cdot \left\{1 - \frac{a}{N}\right\}^{N-n}. \tag{17}$$

In the limit for large N , $H(n; a)$ can be approximated by the use of the approximations

$$\begin{aligned} \binom{N}{n} &\sim \frac{N^n}{n!} \\ \left\{1 - \frac{a}{N}\right\}^{N-n} &\sim e^{-a}, \end{aligned}$$

and we obtain

$$\begin{aligned} H(n, a) &\sim \frac{N^n}{n!} \cdot \frac{a^{n+1}}{N^{n+1}} \cdot e^{-a} \\ &= \frac{a^n}{n!} e^{-a} \cdot \frac{a}{N}. \end{aligned} \tag{18}$$

a is still a discrete variable, but for its larger values it can be replaced by a continuous variable, a , in the following way:

The probability, $H(n; a)da$, of having an activity, a , in the interval between a and $a+da$ is defined by

$$\begin{aligned} H(n; a)da &= \sum_{a' > a}^{a < a+da} \frac{a^n}{n!} e^{-a} \frac{a}{N} \\ &\simeq \frac{a^n}{n!} e^{-a} \cdot \frac{a}{N} \sum_{a' > a}^{a < a+da} 1 \end{aligned} \tag{19}$$

for small da . The number, dN , of values of a between a and $a+da$ is given by (16) to be

$$\sum_{\substack{a < a+da \\ a > a}} 1 \sim \frac{N}{a} da;$$

hence

$$H(n; a)da = \frac{a^n}{n!} e^{-a} da. \quad (20)$$

The expression for $H(n; a)$ looks like the expression for a Poisson distribution indicating the probability, $G(a; n)$, of observing n decays when the activity of the source is a . It should, however, be noted that in the Poisson distribution n is the variable and a is a known constant, while in equation (20) n is the known observable and a the unknown activity in the probability function, $H(n; a)$.

The mean value $E [a]$ of $H(n; a)$ is given by

$$\begin{aligned} E [a] &= \int_0^{\infty} a \cdot \frac{a^n}{n!} e^{-a} da \\ &= \frac{(n+1)!}{n!} \\ &= n+1. \end{aligned} \quad (21)$$

By way of comparison it should be mentioned that the mean value, $E [n]$, of the Poisson distribution

$$G(a; n) = \frac{a^n}{n!} e^{-a}$$

is

$$\begin{aligned} E [n] &= \sum_{n=0}^{\infty} n \cdot \frac{a^n}{n!} e^{-a} \\ &= a. \end{aligned} \quad (22)$$

The difference in the problem manifests itself most explicitly in the cases of $a = 0$ and $n = 0$ expressed in the following way:

If we know that the activity $a = 0$, it follows from the Poisson distribution that the count number is also 0, as

$$g(0, 0) = 1$$

and

$$g(0, n) = 0 \quad (n \geq 1).$$

On the other hand, if we know that no decay has been observed in a time interval t , this does not necessarily mean that the activity is 0. The mean value of the probability function, $H(n; a)$, shows that the mathematical expectation value of the activity a is 1.

It is reasonable to use this mean value (21) as a measure of the unknown activity, a , when n decays have been observed in a given time t . The number of decays per time unit is generally used as a measure of activity (with the unit of 1 curie = $3.7 \cdot 10^{10}$ decays per second), and if n decays have been observed at the time t with a detector of detecting efficiency ϵ , the decay frequency, A_o , is given by

$$A_o = \frac{n+1}{\epsilon \cdot t} . \quad (23)$$

As $\epsilon \leq 1$, $1/t$ is the smallest observable value of A_o .

As a measure of the concentration in probability about the mean value, $E[a]$, we may use the standard deviation, $D[a]$, given by

$$D^2[a] = E[a^2] - E^2[a] , \quad (24)$$

where

$$\begin{aligned} E[a^2] &= \int_0^\infty a^2 \frac{a^n}{n!} e^{-a} da \\ &= (n+2)(n+1). \end{aligned}$$

By applying eq. (21) we obtain

$$D[a] = \sqrt{n+1} . \quad (25)$$

For the sake of curiosity it may be mentioned that the maximum value of the distribution function, $H(n; a)$, is reached for

$$a = n, \quad (26)$$

since

$$\begin{aligned} 0 &= \frac{d}{da} \frac{a^n}{n!} e^{-a} \\ &= \left(\frac{n}{a} - 1\right) \frac{a^n}{n!} e^{-a}. \end{aligned}$$

Derivation of $H(n; a)$ from the Poisson Distribution

In the above analysis, $H(n; a)$ (eq. (20)) appeared as a limit function of the probability function $H_p(n_o; N)$ (eq. (13)) for large values of N and a constant product pN .

In the same limit the probability function, $G_p(N; n_o)$, of the number n_o of decays observed passes into a Poisson distribution:

$$G(a; n) = \frac{a^n}{n!} e^{-a}. \quad (27)$$

Since the Poisson distribution can be derived more generally than as a limit function of a binomial distribution⁴⁾, it may be of interest to obtain the probability distribution, $H(n; a)$, direct on the basis of a Poisson distribution.

The probability of observing n decays in a given time interval is now given by (27), where $a \geq 0$ is a known parameter. The problem is to find the probability, $H(n; a)da$, of the parameter a lying in the interval $(a, a+da)$ when n decays have been observed. For this purpose the probability $Q(n)$ is introduced as the probability of observing n decays independently of the parameter (the activity) a , and the a priori probability, $P(a)da$, that a lies in the interval $(a, a+da)$ independently of an observation of decays.

The simultaneous probability, $S(n, a)da$, is now given by

$$\begin{aligned} S(n, a)da &= P(a)da \cdot G(a; n) \\ &= Q(n) \cdot H(n; a)da. \end{aligned} \quad (28)$$

By integrating these two expressions over all possible values of a we find that

$$\int_0^\infty Q(n) \cdot H(n; a) da = Q(n)$$

$$= \int_0^\infty P(a) \cdot G(a; n) da,$$

which by insertion in (28) gives

$$H(n; a) = \frac{P(a) \cdot G(a; n)}{\int_0^\infty P(a) \cdot G(a; n) da}.$$

Here again it is assumed that all values of $a \geq 0$ have an a priori equal probability by which

$$H(n; a) = \frac{\frac{a^n}{n!} e^{-a}}{\int_0^\infty \frac{a^n}{n!} e^{-a} da}$$

$$= \frac{a^n}{n!} e^{-a}.$$

This result is given by Rainwater and Wu⁵⁾, who do not state the method of derivation.

Repetition of Measurement of a Radioactive Source with Constant Activity

If r measurements are made of the same source with constant decay frequency, A_0 , over equally long time intervals, t , and with constant detector efficiency, ϵ , we find that the probability, $G(a; n_1, \dots, n_r)$, of obtaining a certain set of count numbers (n_1, \dots, n_r) is the product of the probabilities of the individual count numbers. The activity of the source, a , during the time t is given by

$$a = A_0 \cdot \epsilon \cdot t. \quad (29)$$

Here the assumption is made that the individual measurements are independent of each other, which is reasonable since the decay of the individual atoms was assumed to be independent of the decay of the other atoms, and we have

$$\begin{aligned} G(a; n_1, \dots, n_r) &= \frac{a^{n_1}}{n_1!} e^{-a} \cdot \dots \cdot \frac{a^{n_r}}{n_r!} e^{-a} \\ &= \frac{a^{n_1 + \dots + n_r}}{n_1! \cdot \dots \cdot n_r!} e^{-ra}. \end{aligned} \quad (30)$$

The probability, $H(n_1, \dots, n_r; a)da$, of the activity, a , lying in the interval $(a, a+da)$ can now be derived, still on the assumption that all values of $a \geq 0$ are a priori equally possible.

We find

$$\begin{aligned} H(n_1, \dots, n_r; a)da &= \frac{\frac{n_1 + \dots + n_r}{a} \frac{a^{n_1 + \dots + n_r}}{n_1! \cdot \dots \cdot n_r!} e^{-ra} da}{\int_0^\infty \frac{n_1 + \dots + n_r}{a} \frac{a^{n_1 + \dots + n_r}}{n_1! \cdot \dots \cdot n_r!} e^{-ra} da} \\ &= \frac{(r \cdot a)^{n_1 + \dots + n_r}}{(n_1 + \dots + n_r)!} e^{-ra} dra, \end{aligned} \quad (31)$$

which is identical with the probability of having an activity $r \cdot a$ in the interval $(ra, ra+dra)$ when $n_1 + \dots + n_r$ decays have been observed in the time rt .

Equation (31) shows that if a certain time interval, T , has been given for observation of decays in a radioactive source, no principal advantage is obtained by dividing this interval into r subintervals, since the probability distribution of the activity depends only on the total number of decays observed during the total observation time. Therefore it is a moot point whether the phrase "repetition of the measurement" is correct in this case, as it is in reality identical with a prolongation of the measuring time. In practice such a division of the existing time is often used as a means of estimating from the distribution of the individual count numbers whether the efficiency of the detector has stayed constant throughout the experiment.

$H(n; a)$ is Asymptotically Normal

We found above that the distribution function, $H(n; a)$, has the mean value

$$E[a] = n + 1$$

and the standard deviation

$$D[a] = \sqrt{n+1}.$$

The characteristic function, $\varphi(t)$, of the standardized variable (ref. 2),

$$\xi = \frac{a - (n+1)}{\sqrt{n+1}}, \quad (32)$$

is given by

$$\begin{aligned} \varphi(t) &= E \left[e^{it \frac{a - (n+1)}{\sqrt{n+1}}} \right] \\ &= \int_0^\infty e^{it \frac{a - (n+1)}{\sqrt{n+1}}} \frac{a^n}{n!} e^{-a} da \\ &= e^{it \sqrt{n+1}} \left\{ 1 - \frac{it}{\sqrt{n+1}} \right\}^{-(n+1)} \\ &= \left\{ e^{\frac{it}{\sqrt{n+1}}} \left(1 - \frac{it}{\sqrt{n+1}} \right) \right\}^{-(n+1)}, \end{aligned}$$

which for each fixed t we can write as

$$\varphi(t) = \left\{ 1 + \frac{t^2}{2(n+1)} + \mathfrak{d} \frac{t^3}{(n+1)^{3/2}} \right\}^{-(n+1)}$$

when choosing n sufficiently large, and for $|\mathfrak{d}| \leq 1$.

For $n + 1 \rightarrow \infty$, $\varphi(t) \rightarrow e^{-\frac{t^2}{2}}$, so that the distribution function, $H(n; a)$, goes towards a normal distribution with the mean value $n + 1$ and the standard deviation $\sqrt{n+1}$.

The Difference between Two Activities, a_1 and a_2

An interesting case is the one in which observation of decays in two radioactive sources has given the count numbers n_1 and n_2 respectively. The two probability distributions for the activities a_1 and a_2 are given by

$$H(n_1; a_1) \text{ and } H(n_2; a_2)$$

respectively.

If the difference, ζ , between a_1 and a_2 is taken as a new statistical variable, we must consider the probability, $H(n_1, n_2; \zeta) d\zeta$, that $\zeta = a_1 - a_2$ lies in the interval $(\zeta, \zeta + d\zeta)$.

Here a general theorem (appendix 1) dealing with the semi-invariants for a sum of two statistical variables may be used, which says that the n^{th} semi-invariant, $\kappa_{\zeta n}$, for a variable, ζ , given by

$$\zeta = x \cdot a_1 + y \cdot a_2$$

can be written as

$$\kappa_{\zeta n} = x^n \kappa_{a_1 n} + y^n \kappa_{a_2 n},$$

where $\kappa_{a_1 n}$ and $\kappa_{a_2 n}$ are the n^{th} semi-invariants for the distribution functions for a_1 and a_2 respectively; x and y may have any real value.

Further, remembering that

$$\kappa_{\zeta 1} = E[\zeta]$$

and

$$\kappa_{\zeta 2} = D^2[\zeta],$$

we find

$$\begin{aligned}
 E[a_1 - a_2] &= (+1)E[a_1] + (-1)E[a_2] \\
 &= (n_1 + 1) - (n_2 + 1) \\
 &= n_1 - n_2,
 \end{aligned} \tag{33}$$

and

$$\begin{aligned}
 D^2[a_1 - a_2] &= (+1)^2 D^2[a_1] + (-1)^2 D^2[a_2] \\
 &= n_1 + 1 + n_2 + 1 \\
 &= n_1 + n_2 + 2.
 \end{aligned} \tag{34}$$

Background Counts

The background of a detector is the number of registrations which, indistinguishable from registrations of decays in the source, occur without the source being present in the detector. These incorrect registrations have many causes, depending on the construction of the detector and the decays intended for registration. All these causes except contamination of the detector by radioactive material may in practice be eliminated by appropriate construction. To get rid of short-lived radioactive contamination, the detector should be kept isolated for a sufficiently long period, and care should be taken that the placing of the source in the detector does not again cause contamination with short-lived activities. The task under consideration is, therefore, to determine the activity of a source by means of a detector always provided with some active sources, information of which can be obtained only by observing the number of registrations caused by them.

Here it is fortunate that the Poisson distribution is additive, which means that the sum of the count numbers n_1 and n_2 , both of which have been found according to a Poisson distribution, is also Poisson distributed.

Thus, if two activities, α and β , are given, they will cause the count numbers n_1 and n_2 with the probabilities

$$g(\alpha; n_1) = \frac{\alpha^{n_1}}{n_1!} e^{-\alpha}$$

and

$$g(\beta; n_2) = \frac{\beta^{n_2}}{n_2!} e^{-\beta}$$

respectively.

If only the sum, $N = n_1 + n_2$, is observed, the probability of obtaining a certain N is given by

$$\begin{aligned} G(a, \beta; N) &= \sum_{n_1=0}^N \frac{a^{n_1}}{n_1!} e^{-a} \cdot \frac{\beta^{N-n_1}}{(N-n_1)!} e^{-\beta} \\ &= \frac{(a+\beta)^N}{N!} e^{-(a+\beta)} \sum_{n_1=0}^N \frac{N!}{n_1!(N-n_1)!} \left(\frac{a}{a+\beta}\right)^{n_1} \left(\frac{\beta}{a+\beta}\right)^{N-n_1} \\ &= \frac{(a+\beta)^N}{N!} e^{-(a+\beta)}. \end{aligned} \quad (35)$$

By repeated application of this formula (35) the various activities causing the background in a detector can be combined to an activity, b (background activity), which is defined as the fictive activity causing the background. Now an activity, a , is measured with a counter with the background activity b , first by observing the count number n_1 caused by the sum of the activities, $a + b$, secondly by observing the count number n_2 caused by the background activity when the source is not present. To simplify the case it is assumed that the two observation periods are equally long.

The difference, $n_1 - n_2$, between the two count numbers observed is generally used as a measure of the activity a , which, as shown above, is a correct measure of the difference between two Poisson-distributed activities. But in so doing we disregard the extra information, in the case of counting with background, that the difference

$$(a+b) - b = a \quad (36)$$

is positive.

In order that this extra information may be utilized, the probability, $H(n_1, n_2; a)da$, of the activity a lying in the interval $(a, a+da)$ has to be derived on the basis that the count numbers n_1 and n_2 have been observed.

The conditional probability, $G(a, b; n_1)$, of observing n_1 when a and b are given, is found to be

$$G(a, b; n_1) = \sum_{K=0}^{n_1} \frac{a^K}{K!} e^{-a} \frac{b^{n_1-K}}{(n_1-K)!} e^{-b} \quad (37)$$

$$= \frac{(a+b)^{n_1}}{n_1!} e^{-(a+b)}, \quad (38)$$

and the probability, $G(b; n_2)$, of observing n_2

$$G(b; n_2) = \frac{b^{n_2}}{n_2!} e^{-b}. \quad (39)$$

The conditional probability of having a in the interval $(b, b+db)$ when the numbers of counts observed are n_1 and n_2 is denoted

$$H(n_1, n_2; a, b)da db. \quad (40)$$

The unconditional probabilities (a priori probabilities) of having a in the interval $(a, a+da)$ and b in the interval $(b, b+db)$ are denoted

$$P(a)da \quad \text{and} \quad P(b)db \quad (41)$$

respectively.

The unconditional probability of observing a count number n_1 by measurement of source plus background and a count number n_2 by measurement of background only is denoted

$$Q(n_1, n_2). \quad (42)$$

The simultaneous probability, $S(n_1, n_2, a, b)$, of having a in the interval $(a, a+da)$ and b in the interval $(b, b+db)$ while observing n_1 and n_2 can now be written in two ways:

$$\begin{aligned} S(n_1, n_2, a, b)da db &= Q(n_1, n_2) \cdot H(n_1, n_2; a, b)da db \\ &= P(a)da \cdot P(b)db \cdot G(a, b; n_1) \cdot G(b; n_2). \end{aligned} \quad (43)$$

If these two expressions are integrated over the definition domains of a and b ($a \geq 0$ and $b \geq 0$), we obtain

$$\begin{aligned}
 Q(n_1, n_2) \int_0^\infty \int_0^\infty H(n_1, n_2; a, b) da db &= Q(n_1, n_2) \\
 &= \int_0^\infty \int_0^\infty P(a) P(b) G(a, b; n_1) G(b; n_2) da db. \quad (44)
 \end{aligned}$$

In order to be a probability distribution, $H(n_1, n_2; a, b)da db$ must be normalized. Now it is possible to give an expression for this probability; we find

$$H(n_1, n_2; a, b) = \frac{P(a) P(b) G(a, b; n_1) G(b; n_2)}{\int_0^\infty \int_0^\infty P(a) P(b) G(a, b; n_1) G(b; n_2) da db}. \quad (45)$$

Here again it is reasonable to assume a priori that every positive value of a and b is equally probable, whereby we obtain

$$\begin{aligned}
 H(n_1, n_2; a, b) &= \frac{G(a, b; n_1) G(b; n_2)}{\int_0^\infty \int_0^\infty G(a, b; n_1) G(b; n_2) da db} \\
 &= \frac{\frac{(a+b)^{n_1}}{n_1!} e^{-(a+b)} \frac{b^{n_2}}{n_2!} e^{-b}}{\int_0^\infty \int_0^\infty \frac{(a+b)^{n_1}}{n_1!} e^{-(a+b)} \cdot \frac{b^{n_2}}{n_2!} e^{-b} da db}. \quad (46)
 \end{aligned}$$

Calculating the denominator in (46) by using expression (37) for $G(a, b; n_1)$, we have

$$\begin{aligned}
 &\int_0^\infty \int_0^\infty \frac{(a+b)^{n_1}}{n_1!} e^{-(a+b)} \frac{b^{n_2}}{n_2!} e^{-b} da db \\
 &= \sum_{K=0}^{n_1} \int_0^\infty \frac{a^K}{K!} e^{-a} da \int_0^\infty \frac{b^{n_1-K+n_2}}{(n_1-K)! n_2!} e^{-2b} db \\
 &= \sum_{K=0}^{n_1} \frac{(n_1-K+n_2)!}{(n_1-K)! n_2! 2^{n_1-K+n_2+1}}.
 \end{aligned} \quad (47)$$

Since only the magnitude of a is of general interest, we integrate $H(n_1, n_2; a, b)$ over all values of $b \geq 0$, thus obtaining the conditional probability distribution, $H(n_1, n_2; a)$, for a based on the observation of the count numbers n_1 and n_2 :

$$\begin{aligned}
 H(n_1, n_2; a) &= \int_0^\infty H(n_1, n_2; a, b) db \\
 &= \frac{\sum_{K=0}^{n_1} \frac{a^K}{K!} e^{-a} \frac{(n_1 - K + n_2)!}{(n_1 - K)! n_2! 2^{n_1 - K + n_2 + 1}}}{\sum_{K=0}^{n_1} \frac{(n_1 - K + n_2)!}{(n_1 - K)! n_2! 2^{n_1 - K + n_2 + 1}}} \\
 &= \frac{\sum_{K=0}^{n_1} \frac{a^K}{K!} e^{-a} \frac{(n_1 - K + n_2)!}{(n_1 - K)!} 2^K}{\sum_{K=0}^{n_1} \frac{(n_1 - K + n_2)!}{(n_1 - K)!} 2^K}. \tag{48}
 \end{aligned}$$

Mean Value and Standard Deviation in the Case with a Background

The mean value, $E[a]$, of the probability distribution $H(n_1, n_2; a)$ is found to be

$$\begin{aligned}
 E[a] &= \int_0^\infty a \cdot H(n_1, n_2; a) da \\
 &= \frac{\sum_{K=0}^{n_1} \int_0^\infty \frac{a^{K+1}}{K!} e^{-a} da \cdot \frac{(n_1 - K + n_2)! 2^K}{(n_1 - K)!}}{\sum_{K=0}^{n_1} \frac{(n_1 - K + n_2)! 2^K}{(n_1 - K)!}} \\
 &= \frac{\sum_{K=0}^{n_1} (K+1) \frac{(n_1 - K + n_2)! 2^K}{(n_1 - K)!}}{\sum_{K=0}^{n_1} \frac{(n_1 - K + n_2)! 2^K}{(n_1 - K)!}}. \tag{49}
 \end{aligned}$$

The expression for $E[a]$ can be calculated on a digital computer such as DASK, and in table 1 $E[a]$ is given for values of n_1 and n_2 up to 51.

It is desirable, however, to make a certain analysis of the expression and especially to investigate the limit for large values of n_1 .

Simplified, the expression can be written

$$E[a] = \frac{\sum_{K=0}^{n_1} (K+1) f(K)}{\sum_{K=0}^{n_1} f(K)}, \quad (50)$$

where $f(K) > 0$, and as the addends in the numerator are multiplied by a factor $(K+1) \geq 1$, it is obvious that

$$E[a] \geq 1, \quad (51)$$

where the sign of equality holds only for $n_1 = 0$ independently of n_2 .

To investigate $E[a]$ for large values of n_1 with constant n_2 we substitute r for $n_1 - K$ in the expression (49), and obtain

$$E[a] = n_1 - n_2 + \frac{\frac{(n_1+n_2+1)!}{n_1}}{\sum_{r=0}^{n_1} \frac{n_1! 2^r}{(r+n_2)! r! 2^r}}, \quad (52)$$

and for each constant n_2 it can be found (appendix 2) that

$$\frac{\frac{(n_1+n_2+1)!}{n_1}}{\sum_{r=0}^{n_1} \frac{n_1! 2^r}{(r+n_2)! r! 2^r}} \rightarrow 0 \quad \text{for } n_1 \rightarrow \infty, \quad (53)$$

from which the measure generally used for the source strength a is obtained, since then

$$E[a] \rightarrow n_1 - n_2. \quad (54)$$

The standard deviation, $D[a]$, given by

$$D^2[a] = E[a^2] - E^2[a]$$

is used as a measure of the concentration in probability about the mean value, $E[a]$, $E[a^2]$ being found by means of the expression

$$\begin{aligned} E[a^2] &= \int_0^\infty a^2 H(n_1, n_2; a) da \\ &= \frac{\sum_{K=0}^{n_1} \int_0^\infty \frac{a^{K+2}}{K!} e^{-a} da \frac{(n_1-K+n_2)! 2^K}{(n_1-K)!}}{\sum_{K=0}^{n_1} \frac{(n_1-K+n_2)! 2^K}{(n_1-K)!}} \\ &= \frac{\sum_{K=0}^{n_1} (K+1)(K+2) \frac{(n_1-K+n_2)! 2^K}{(n_1-K)!}}{\sum_{K=0}^{n_1} \frac{(n_1-K+n_2)! 2^K}{(n_1-K)!}}. \end{aligned} \quad (55)$$

From this, $D^2[a]$ is found (appendix 2) to be

$$\begin{aligned} D^2[a] &= n_1 + n_2 + 3 - (n_1 - n_2) \cdot \frac{\frac{(n_1+n_2+1)!}{n_1! 2^n_1}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}} \\ &\quad - \left\{ \frac{\frac{(n_1+n_2+1)!}{n_1! 2^n_1}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}} \right\}^2 \end{aligned} \quad (56)$$

$\Rightarrow n_1 + n_2 + 3$ for $n_1 \rightarrow \infty$ and constant n_2 .

Table 1 also gives values of the relative standard deviation $D[a]/E[a]$ for n_1 and n_2 up to 51.

In the limit $n_1 \rightarrow 0$ it is seen from formula (56) that

$$\lim_{n_1 \rightarrow 0} D^2[a] = n_2 + 3 - (-n_2) \cdot (n_2+1) - (n_2+1)^2 \\ = 1,$$

and it has already been shown (formula 50) that

$$\lim_{n_1 \rightarrow 0} E[a] = 1,$$

which means that the relative standard deviation

$$\frac{D[a]}{E[a]} \rightarrow 1 \text{ when } n_1 \rightarrow 0 \quad (57)$$

independently of the value of n_2 .

Approximative Expressions for $E[a]$ and $D[a]/E[a]$

The mean value $E[a]$ of the probability distribution for the activity a was given by equation (52):

$$E[a] = n_1 - n_2 + \sum_{r=0}^{n_1} \frac{\frac{(n_1+n_2+1)!}{n_1! 2^r}}{\frac{(r+n_2)!}{r! 2^r}}, \quad (52)$$

which was the exact expression on the basis of the theory given.

In order to derive a more convenient formula we transform the denominator (appendix 3) into

$$\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r} = n_2! 2^{n_2+1} \cdot \int_0^{\frac{1}{2}} \frac{(n_1+n_2+1)!}{n_1! n_2!} x^{n_2} (1-x)^{n_1} dx, \quad (58)$$

where the integrand

$$\frac{(n_1+n_2+1)!}{n_1! n_2!} x^{n_2} (1-x)^{n_1} = \beta(p, q; x)$$

is the Beta-distribution (ref. 2) with the parameters

$$p = n_2 + 1 \quad \text{and} \quad q = n_1 + 1.$$

For larger values of p and q this distribution function tends to the normal distribution (ref. 2), which can then be used in deriving an approximation, since

$$\begin{aligned} E[a] &= n_1 - n_2 + \frac{\frac{1}{2} \frac{(n_1+n_2+1)!}{n_1! n_2!} (\frac{1}{2})^{n_1} (\frac{1}{2})^{n_2}}{\int_0^{\frac{1}{2}} \frac{(n_1+n_2+1)!}{n_1! n_2!} x^{n_2} (1-x)^{n_1} dx} \\ &\sim n_1 - n_2 + \frac{\frac{1}{2\sigma} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}}{\int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt}, \end{aligned} \quad (59)$$

where

$$\sigma^2 = \frac{(n_1+1)(n_2+1)}{(n_1+n_2+2)^2(n_1+n_2+3)} \quad (60)$$

and

$$x = \frac{1}{\sigma} \left(\frac{1}{2} - \frac{n_2+1}{n_1+n_2+2} \right). \quad (61)$$

The approximation is only of interest when n_1 and n_2 are fairly close to each other, and then

$$\begin{aligned} \frac{1}{\sigma^2} &= \left(1 + \frac{n_2+1}{n_1+1}\right) \left(1 + \frac{n_1+1}{n_2+1}\right) (n_1+n_2+3) \\ &= (2 + y + \frac{1}{y})(n_1+n_2+3) \\ &\sim 4(n_1+n_2+3), \end{aligned} \quad (62)$$

where

$$y = \frac{n_2 + 1}{n_1 + 1}.$$

The error introduced by this approximation is seen to be of second order in y .

Now the final approximation for the mean value of the activity a is found to be

$$E[a] \sim n_1 - n_2 + \sqrt{n_1 + n_2 + 3} \cdot \kappa(x), \quad (63)$$

where

$$x = \frac{n_1 - n_2}{\sqrt{n_1 + n_2 + 3}} \quad (64)$$

and

$$\kappa(x) = \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}}{\int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt}. \quad (65)$$

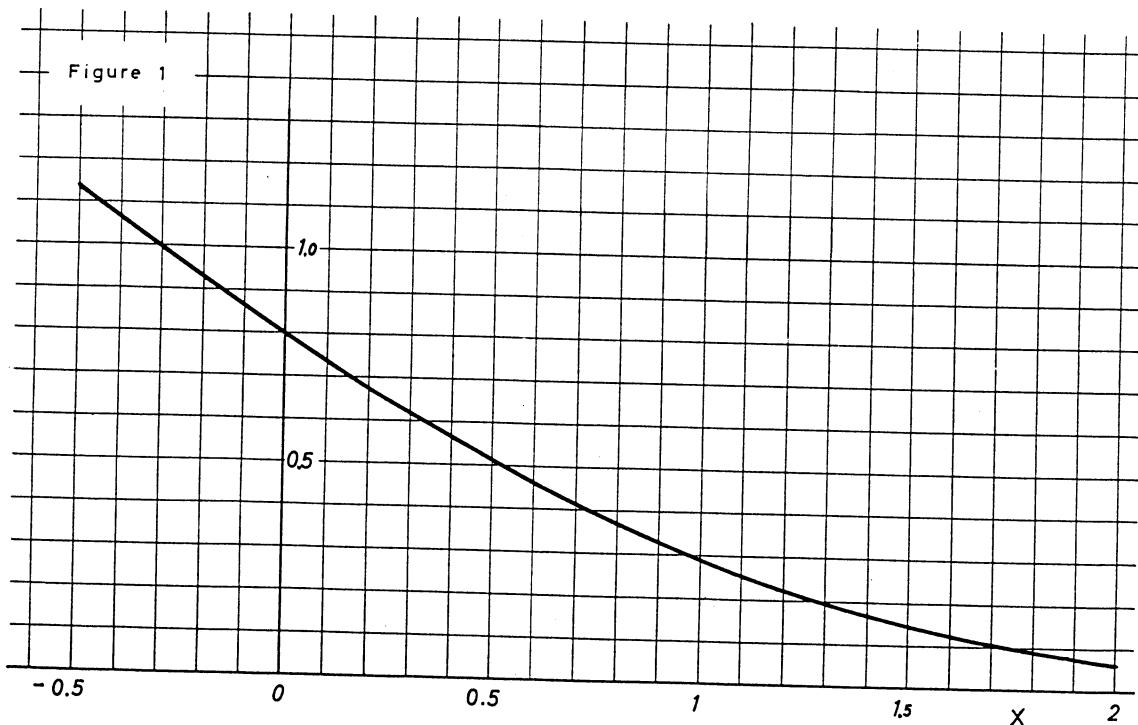


Figure 1, calculated on the basis of a standard table of the error function (e.g. ref. 6), shows the value of $\kappa(x)$ for the range $-0.5 < x < 2.0$.
The formula for the squared standard deviation,

$$D^2[a] = n_1 + n_2 + 3 - (n_1 - n_2) \frac{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}}{\frac{(n_1+n_2+1)!}{n_1! 2^{n_1}}} - \left\{ \frac{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}}{\frac{(n_1+n_2+1)!}{n_1! 2^{n_1}}} \right\}^2, \quad (66)$$

is found in appendix 2. The introduction of the same approximations as those used for $E[a]$ will give

$$D^2[a] \sim n_1 + n_2 + 3 - (n_1 - n_2) \sqrt{n_1 + n_2 + 3} \kappa(x) - (n_1 + n_2 + 3) \kappa^2(x), \quad (67)$$

and by division by the expression for $E[a]$ we obtain

$$\frac{D[a]}{E[a]} \sim \frac{\sqrt{1-x \kappa(x) - \kappa^2(x)}}{x + \kappa(x)}. \quad (68)$$

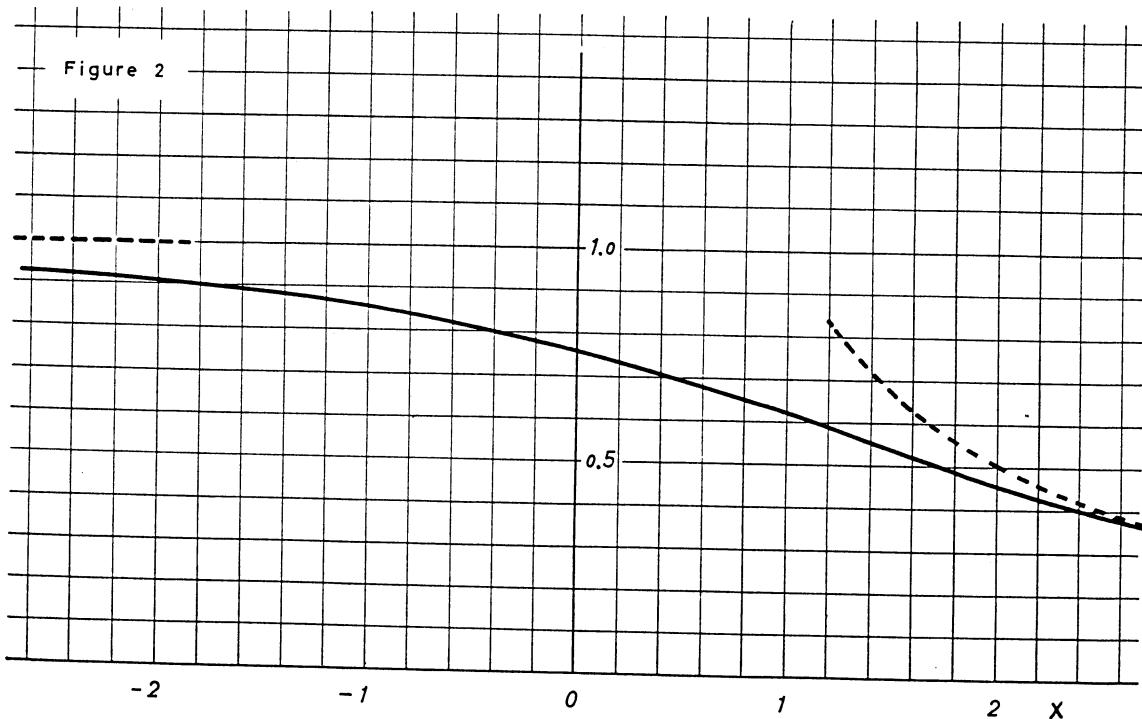


Figure 2, together with formula (68), shows the function x^{-1} (dashed line) since for constant n_2

$$\frac{D[a]}{E[a]} \rightarrow \frac{\sqrt{n_1+n_2+3}}{n_1-n_2} = \frac{1}{x} \text{ as } n_1 \rightarrow \infty. \quad (69)$$

Summary

After an analysis of a set of assumptions which leads to the Poisson distribution for the observation of radioactive decays, the problem of the magnitude of the apparent activity of the radioactive source is considered.

It is assumed that the decays in the source in question are observed by means of a detector that exhibits a certain background count rate. During equal periods of observation the radioactive source plus the background are assumed to have given the number of counts n_1 and the background alone the number n_2 . On this basis the probability,

$$H(n_1, n_2; a)da,$$

of the unknown activity a_1 lying in the interval $(a, a+da)$ is derived (formula 48):

$$H(n_1, n_2; a) = \frac{\sum_{K=0}^{n_1} \frac{a^K}{K!} e^{-a} \frac{(n_1-K+n_2)!}{(n_1-K)!} 2^K}{\sum_{K=0}^{n_1} \frac{(n_1-K+n_2)!}{(n_1-K)!} 2^K}.$$

It is further shown that no extra information is gained on the magnitude of the activity by a division of the observation time.

In order to find a representative measure of the activity, the mean value $E[a]$ and the standard deviation $D[a]$ of the above-mentioned probability function are derived (formulae 52 and 56 respectively):

$$E[a] = n_1 - n_2 + \frac{\frac{(n_1+n_2+1)!}{n_1! 2^{n_1}}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}}$$

and

$$D^2 [a] = n_1 + n_2 + 3 - (n_1 - n_2) \cdot \frac{\sum_{r=0}^{n_1} \frac{(n_1 + n_2 + 1)!}{n_1! 2^r}}{\left\{ \frac{(n_1 + n_2 + 1)!}{n_1! 2^{n_1}} \right\}^2}$$

Since these formulae are rather inconvenient even for very small count numbers, a tabulation of $E[a]$ and $D[a]/E[a]$ is made for n_1 and n_2 , both up to 51 (table 1).

Approximative expressions are derived, which give the mean value of the activity

$$E[a] \sim n_1 - n_2 + \sqrt{n_1 + n_2 + 3} \cdot \kappa(x),$$

where

$$x = \frac{n_1 - n_2}{\sqrt{n_1 + n_2 + 3}}$$

and

$$\kappa(x) = \frac{\frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}}{\int_{-\infty}^x \frac{1}{\sqrt{2\pi}} e^{-\frac{t^2}{2}} dt}.$$

The relative standard deviation is found to be approximated by

$$\frac{D[a]}{E[a]} \sim \frac{\sqrt{1-x \kappa(x) - \kappa^2(x)}}{x + \kappa(x)}.$$

$\kappa(x)$ can be taken from fig. 1 and $D[a]/E[a]$ from fig. 2.

When $n_1 \gg n_2$ in such a way that x is large, it is found that

$$E[a] \rightarrow n_1 - n_2$$

and

$$\frac{D[a]}{E[a]} \rightarrow \frac{1}{x} = \frac{\sqrt{n_1 + n_2 + 3}}{n_1 - n_2}.$$

Acknowledgements

My special thanks are due to Professor Otto Kofoed-Hansen, whose inquisitorial questions considerably increased the stringency of the argumentation; further to Peter Villemoes and Helge Vilstrup, who prepared the ALGOL programme for the calculation of table 1, and to Leif Løvborg for checking the manuscript.

The report has been carefully typed by Mrs. Karna Hansen, the figures have been drawn by Kurt Holmqvist, and special signs in the text have been written by Knud J. Nielsen.

The heavy burden of keeping the text in the vicinity of correct English grammar has been carried bravely by Flemming Steenbuch.

Appendix 1Theorem on Semi-invariants

If a statistical variable, ξ , is given as a weighed sum of two independent statistical variables, η and ζ , by

$$\xi = a \cdot \eta + b \zeta,$$

where a and b are real numbers, the n^{th} semi-invariant $\kappa_{n,\xi}$ of ξ is

$$\kappa_{n,\xi} = a^n \kappa_{n,\eta} + b^n \kappa_{n,\zeta},$$

where $\kappa_{n,\eta}$ and $\kappa_{n,\zeta}$ are the n^{th} semi-invariants of η and ζ respectively.

Proof:

The characteristic function, $\varphi_\xi(t)$, of the statistical variable ξ (ref. 2) is expressed by

$$\varphi_\xi(t) = \varphi_\eta(t) \cdot \varphi_\zeta(t)$$

since η and ζ were independent variables. A power-series expansion of the logarithm of the characteristic function yields

$$\begin{aligned} \log \varphi_\xi(t) &= \sum_{n=1}^k \frac{\kappa_{n,\xi}}{n!} (it)^n + \sigma(t^k) \\ &= \log \varphi_\eta(t) + \log \varphi_\zeta(t) \\ &= \sum_{n=1}^k \frac{\kappa_{n,\eta}}{n!} (iat)^n + \sigma(t^k) + \sum_{n=1}^k \frac{\kappa_{n,\zeta}}{n!} (ibt)^n + \sigma(t^k). \end{aligned}$$

Hence

$$\sum_{n=1}^k \frac{\kappa_{n,\xi}}{n!} (it)^n + \sigma(t^k) = \sum_{n=1}^k \left(\frac{a^n \kappa_{n,\eta} + b^n \kappa_{n,\zeta}}{n!} \right) (it)^n + \sigma(t^k).$$

The identity of the coefficients to the same power of the variable immediately yields the theorem.

Appendix 2Calculation of Expressions for E [a] and D [a]

The expression for the mean value of the activity as given by formula (49),

$$E [a] = \frac{\sum_{K=0}^{n_1} (K+1) \frac{(n_1-K+n_2)! 2^K}{(n_1-K)!}}{\sum_{K=0}^{n_1} \frac{(n_1-K+n_2)! 2^K}{(n_1-K)!}}, \quad (49)$$

is transformed by the substitution $r = n_1 - K$ into

$$E [a] = \frac{\sum_{r=0}^{n_1} (n_1+1-r) \frac{(r+n_2)!}{r!} 2^{n_1-r}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r!} 2^{n_1-r}}$$

$$= n_1 + 1 - \frac{\sum_{r=0}^{n_1} r \frac{(r+n_2)!}{r! 2^r}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}}$$

$$= n_1 - n_2 + \frac{\frac{(n_1+n_2+1)!}{n_1}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}}.$$

In the last step we use the relation

$$\sum_{r=0}^{n_1} r \frac{(r+n_2)!}{r! 2^r} = \sum_{r=1}^{n_1} \frac{(r+n_2)(r-1+n_2)!}{(r-1)! 2^{r-1} \cdot 2}$$

$$= \frac{1}{2} \sum_{r=1}^{n_1} \frac{(r-1+n_2+1)(r-1+n_2)!}{(r-1)! 2^{r-1}}$$

$$= \frac{1}{2} \left\{ \sum_{r=1}^{n_1} (r-1) \frac{(r-1+n_2)!}{(r-1)! 2^{r-1}} + (n_2+1) \sum_{r=1}^{n_1} \frac{(r-1+n_2)!}{(r-1)! 2^{r-1}} \right\}$$

$$= \frac{1}{2} \left\{ \sum_{r=0}^{n_1-1} r \frac{(r+n_2)!}{r! 2^r} + (n_2+1) \sum_{r=0}^{n_1-1} \frac{(r+n_2)!}{r! 2^r} \right\}$$

$$= \frac{1}{2} \left\{ \sum_{r=0}^{n_1} r \frac{(r+n_2)!}{r! 2^r} + (n_2+1) \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r} \right.$$

$$\left. - n_1 \frac{(n_1+n_2)!}{n_1! 2^{n_1}} - (n_2+1) \frac{(n_1+n_2)!}{n_1! 2^{n_1}} \right\}$$

$$= (n_2+1) \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r} - \frac{(n_1+n_2+1)!}{n_1! 2^{n_1}} .$$

The preceding calculations are somewhat facilitated by the use of the abbreviations

$$X = \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}$$

and

$$Y = \frac{(n_1+n_2+1)!}{n_1! 2^{n_1}} ,$$

by which

$$E[a] = n_1 - n_2 + \frac{Y}{X}.$$

In order to show that

$$\frac{Y}{X} = \frac{\frac{(n_1+n_2+1)!}{n_1! 2^{n_1}}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}} \rightarrow 0 \quad \text{for } n_1 \rightarrow \infty \text{ and } n_2 \text{ constant}$$

we first show that the denominator is non-vanishing. Since the denominator is a sum of positive elements, it is bigger than e.g. the first of the addends:

$$\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r} \gg \frac{(0+n_2)!}{0! 2^0} = n_2!,$$

which is non-vanishing when $n_2 \geq 1$.

Using Sterling's formula

$$n! \sim e^{-n} \cdot n^n \cdot \sqrt{2\pi n},$$

we then see that the numerator tends to zero for $n_1 \rightarrow \infty$ and n_2 constant, since

$$\frac{(n_1+n_2+1)!}{n_1! 2^{n_1}} \sim \frac{e^{-(n_1+n_2+1)} (n_1+n_2+1)^{(n_1+n_2+1)}}{e^{-n_1} n_1^{n_1} \sqrt{2\pi n_1} 2^{n_1}} \sqrt{2\pi (n_1+n_2+1)}$$

$$= \frac{1}{e^{n_2+1}} \left(\left(1 + \frac{n_2+1}{n_1} \right)^{\frac{n_1}{n_2+1} + 1} \right)^{n_2+1} \frac{n_1^{n_2+1}}{2^{n_1}} \sqrt{1 + \frac{n_2+1}{n_1}}$$

$$\frac{(n_1 + n_2 + 1)!}{n_1! 2^{n_1}} \rightarrow \frac{1}{e^{n_2+1}} \cdot e^{n_2+1} \cdot 0 \cdot 1 \\ = 0$$

for $n_1 \rightarrow \infty$ and n_2 constant, which concludes the proof.

By substituting r for $n_1 - K$ we transform the expression

$$E[a^2] = \frac{\sum_{K=0}^{n_1} (K+1)(K+2) \frac{(n_1-K+n_2)! 2^K}{(n_1-K)!}}{\sum_{K=0}^{n_1} \frac{(n_1-K+n_2)! 2^K}{(n_1-K)!}}$$

into

$$E[a^2] = \frac{\sum_{r=0}^{n_1} (n_1+1-r)(n_1+2-r) \frac{(r+n_2)! 2^{n_1-r}}{r!}}{\sum_{r=0}^{n_1} \frac{(r+n_2)! 2^{n_1-r}}{r!}}$$

$$= (n_1+1)(n_1+2) - (2n_1+3) \frac{\sum_{r=0}^{n_1} r \frac{(r+n_2)!}{r! 2^r}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}}$$

$$+ \frac{\sum_{r=0}^{n_1} r^2 \frac{(r+n_2)!}{r! 2^r}}{\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r}} .$$

Here we find that

$$\sum_{r=0}^{n_1} r^2 \frac{(r+n_2)!}{r! 2^r} = \sum_{r=1}^{n_1} r \frac{(r+n_2)!}{(r-1)! 2^r}$$

$$= \sum_{r=1}^{n_1} (r-1+1)(r-1+n_2+1) \frac{(r-1+n_2)!}{(r-1)! 2^{r-1} \cdot 2}$$

$$\begin{aligned}
&= \frac{1}{2} \left\{ \sum_{r=1}^{n_1} (r-1)^2 \frac{(r-1+n_2)!}{(r-1)! 2^{r-1}} + (n_2+2) \sum_{r=1}^{n_1} (r-1) \frac{(r-1+n_2)!}{(r-1)! 2^{r-1}} \right. \\
&\quad \left. + (n_2+1) \sum_{r=1}^{n_1} \frac{(r-1+n_2)!}{(r-1)! 2^{r-1}} \right\} \\
&= \frac{1}{2} \left\{ \sum_{r=0}^{n_1} r^2 \frac{(r+n_2)!}{r! 2^r} + (n_2+2) \sum_{r=0}^{n_1} r \frac{(r+n_2)!}{r! 2^r} + (n_2+1) \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r} \right. \\
&\quad \left. - n_1^2 \frac{(n_1+n_2)!}{n_1! 2^{n_1}} - (n_2+2) \cdot n_1 \frac{(n_1+n_2)!}{n_1! 2^{n_1}} - (n_2+1) \frac{(n_1+n_2)!}{n_1! 2^{n_1}} \right\} \\
&= (n_2+2) \sum_{r=0}^{n_1} r \frac{(r+n_2)!}{r! 2^r} + (n_2+1) \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r! 2^r} - (n_1+1) \frac{(n_1+n_2+1)!}{n_1! 2^{n_1}} \\
&= (n_2+2) (n_2+1)X - (n_1+n_2+3)Y \\
&= (n_2+3) (n_2+1)X - (n_1+n_2+3)Y.
\end{aligned}$$

The mean value of the squared activity is now found to be

$$\begin{aligned}
E[a^2] &= (n_1+1)(n_1+2) - (2n_1+3) \frac{(n_2+1)X-Y}{X} + \frac{(n_2+3)(n_2+1)X-(n_1+n_2+3)Y}{X} \\
&= (n_1-n_2)^2 + (n_1+n_2+3) + (n_1-n_2) \frac{Y}{X}.
\end{aligned}$$

The square of the standard deviation is then

$$\begin{aligned}
D^2[a] &= E[a^2] - E^2[a] \\
&= (n_1-n_2)^2 + (n_1+n_2+3) + (n_1-n_2) \frac{Y}{X} \\
&\quad - (n_1-n_2 + \frac{Y}{X})^2 \\
&= n_1+n_2+3 - (n_1-n_2) \cdot \frac{Y}{X} - (\frac{Y}{X})^2.
\end{aligned}$$

As shown above,

$$\frac{Y}{X} \rightarrow 0 \text{ for } n_1 \rightarrow \infty \text{ and } n_2 \text{ constant,}$$

and thus

$$D^2 [a] \rightarrow n_1 + n_2 + 3$$

under the same conditions.

Appendix 3

$$\text{Summation of } \sum_{r=0}^{n_1} \frac{(n_2+r)!}{r!} \left(\frac{1}{2}\right)^r$$

The sum occurring in the denominator in the formula for the mean value and the standard deviation is a little unhandy for direct use when the observed count numbers n_1 and n_2 become larger. In order to find an approximative expression we convert the series into a function $f(n_1, n_2, x)$ defined by

$$f(n_1, n_2, x) = \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r!} x^r.$$

This function is differentiable with the differential quotient

$$\begin{aligned} \frac{\partial}{\partial x} f(n_1, n_2, x) &= \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r!} r x^{r-1} \\ &= \sum_{r-1=0}^{n_1-1} \frac{(r-1+n_2+1)!}{(r-1)!} x^{r-1} \\ &= f(n_1-1, n_2+1, x). \end{aligned}$$

Further, we find that

$$\begin{aligned}
 f(n_1, n_2+1, x) &= \sum_{r=0}^{n_1} \frac{(r+n_2+1)!}{r!} x^r \\
 &= x \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r!} r x^{r-1} + (n_2+1) \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r!} x^r \\
 &= x \cdot f(n_1-1, n_2+1, x) + (n_2+1) f(n_1, n_2, x)
 \end{aligned}$$

and

$$\begin{aligned}
 f(n_1, n_2+1, x) &= \sum_{r=0}^{n_1} \frac{(r+n_2+1)!}{r!} x^r \\
 &= \sum_{r=0}^{n_1} \frac{(r+n_2+1)!}{r!} x^r + \frac{(n_1+n_2+1)!}{n_1!} x^{n_1} \\
 &= f(n_1-1, n_2+1, x) + \frac{(n_1+n_2+1)!}{n_1!} x^{n_1}.
 \end{aligned}$$

A combination of these two equations gives

$$\begin{aligned}
 f(n_1-1, n_2+1, x) &= \frac{1}{1-x} (n_2+1) f(n_1, n_2, x) - \frac{(n_1+n_2+1)!}{n_1!} x^{n_1} \\
 &= \frac{\partial}{\partial x} f(n_1, n_2, x).
 \end{aligned}$$

This is a linear first-order differential equation:

$$\frac{\partial}{\partial x} f(n_1, n_2, x) - \frac{n_2+1}{1-x} f(n_1, n_2, x) = - \frac{(n_1+n_2+1)!}{n_1!} \frac{x^{n_1}}{1-x},$$

of the form

$$\frac{dy}{dx} + f(x)y = g(x),$$

where

$$f(x) = - \frac{n_2+1}{1-x}$$

and down above.

$$g(x) = - \frac{(n_1 + n_2 + 1)!}{n_1!} \frac{x^{n_1}}{1-x}.$$

The solution to such an equation is given by

$$y = e^{- \int f(x) dx} \cdot \left\{ \int g(x) e^{\int f(x) dx} dx + C \right\},$$

where

$$\int_0^x f(x) dx = \int_0^x -\frac{n_2 + 1}{1-x} dx$$

$$= \ln(1-x)^{n_2 + 1}$$

and

$$\begin{aligned} \int_0^x g(x) e^{\int_0^x f(x) dx} dx &= \int_0^x -\frac{(n_1 + n_2 + 1)!}{n_1!} \frac{x^{n_1}}{1-x} \cdot (1-x)^{n_2 + 1} dx \\ &= -\frac{(n_1 + n_2 + 1)!}{n_1!} \int_0^x x^{n_1} (1-x)^{n_2} dx. \end{aligned}$$

By insertion we obtain

$$f(n_1, n_2, x) = \frac{1}{(1-x)^{n_2 + 1}} \left\{ -\frac{(n_1 + n_2 + 1)!}{n_1!} \int_0^x x^{n_1} (1-x)^{n_2} dx + C \right\}$$

$$\rightarrow C \quad \text{as } x \rightarrow 0.$$

It is directly seen that

$$f(n_1, n_2, x) = \sum_{r=0}^{n_1} \frac{(r+n_2)!}{r!} x^r$$

$$\rightarrow n_2! \quad \text{as } x \rightarrow 0,$$

and this gives

$$C = n_2! .$$

The solution is thus

$$f(n_1, n_2, x) = \frac{1}{(1-x)^{n_2+1}} \left\{ -\frac{(n_1+n_2+1)!}{n_1!} \int_0^x x^{n_1} (1-x)^{n_2} dx + n_2! \right\} ,$$

and, taking $x = \frac{1}{2}$, we find

$$\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r!} \left(\frac{1}{2}\right)^r = 2^{n_2+1} \left\{ n_2! - \frac{(n_1+n_2+1)!}{n_1!} \int_0^{\frac{1}{2}} x^{n_1} (1-x)^{n_2} dx \right\} .$$

By using the formula ⁶⁾

$$\begin{aligned} \int_0^1 x^{n_1} (1-x)^{n_2} dx &= \int_0^{\frac{1}{2}} x^{n_1} (1-x)^{n_2} dx + \int_{\frac{1}{2}}^1 x^{n_1} (1-x)^{n_2} dx \\ &= \int_0^{\frac{1}{2}} x^{n_1} (1-x)^{n_2} dx + \int_0^{\frac{1}{2}} x^{n_2} (1-x)^{n_1} dx \\ &= \frac{n_1! n_2!}{(n_1+n_2+1)!} \end{aligned}$$

we obtain

$$\sum_{r=0}^{n_1} \frac{(r+n_2)!}{r!} \left(\frac{1}{2}\right)^r = 2^{n_2+1} \frac{(n_1+n_2+1)!}{n_1!} \int_0^{\frac{1}{2}} x^{n_2} (1-x)^{n_1} dx.$$

Table 1

Table 1 gives the mean values $E [a]$ and the relative standard deviations $D [a] / E [a]$ for

$$1 \leq n_1 \leq 51$$

and

$$1 \leq n_2 \leq 51.$$

The table has been calculated on the basis of the exact formulae.

Table 1

n1	n2 = 1	n2 = 2	n2 = 3			
	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.500	8.819 ₁₀ -1	1.400	9.147 ₁₀ -1	1.333	9.354 ₁₀ -1
2	2.091	7.886 ₁₀ -1	1.875	8.406 ₁₀ -1	1.727	8.760 ₁₀ -1
3	2.769	7.104 ₁₀ -1	2.429	7.737 ₁₀ -1	2.188	8.197 ₁₀ -1
4	3.526	6.432 ₁₀ -1	3.061	7.124 ₁₀ -1	2.718	7.658 ₁₀ -1
5	4.350	5.851 ₁₀ -1	3.767	6.562 ₁₀ -1	3.319	7.142 ₁₀ -1
6	5.227	5.349 ₁₀ -1	4.541	6.049 ₁₀ -1	3.991	6.652 ₁₀ -1
7	6.143	4.915 ₁₀ -1	5.372	5.585 ₁₀ -1	4.727	6.191 ₁₀ -1
8	7.089	4.543 ₁₀ -1	6.250	5.169 ₁₀ -1	5.521	5.762 ₁₀ -1
9	8.054	4.222 ₁₀ -1	7.164	4.799 ₁₀ -1	6.366	5.366 ₁₀ -1
10	9.032	3.947 ₁₀ -1	8.106	4.472 ₁₀ -1	7.252	5.006 ₁₀ -1
11	1.002 ₁₀ 1	3.709 ₁₀ -1	9.067	4.184 ₁₀ -1	8.170	4.679 ₁₀ -1
12	1.101 ₁₀ 1	3.503 ₁₀ -1	1.004 ₁₀ 1	3.931 ₁₀ -1	9.112	4.387 ₁₀ -1
13	1.201 ₁₀ 1	3.324 ₁₀ -1	1.103 ₁₀ 1	3.708 ₁₀ -1	1.007 ₁₀ 1	4.125 ₁₀ -1
14	1.300 ₁₀ 1	3.166 ₁₀ -1	1.202 ₁₀ 1	3.513 ₁₀ -1	1.105 ₁₀ 1	3.892 ₁₀ -1
15	1.400 ₁₀ 1	3.028 ₁₀ -1	1.301 ₁₀ 1	3.340 ₁₀ -1	1.203 ₁₀ 1	3.684 ₁₀ -1
16	1.500 ₁₀ 1	2.904 ₁₀ -1	1.401 ₁₀ 1	3.187 ₁₀ -1	1.302 ₁₀ 1	3.500 ₁₀ -1
17	1.600 ₁₀ 1	2.794 ₁₀ -1	1.500 ₁₀ 1	3.051 ₁₀ -1	1.401 ₁₀ 1	3.335 ₁₀ -1
18	1.700 ₁₀ 1	2.695 ₁₀ -1	1.600 ₁₀ 1	2.929 ₁₀ -1	1.501 ₁₀ 1	3.188 ₁₀ -1
19	1.800 ₁₀ 1	2.606 ₁₀ -1	1.700 ₁₀ 1	2.820 ₁₀ -1	1.600 ₁₀ 1	3.057 ₁₀ -1
20	1.900 ₁₀ 1	2.524 ₁₀ -1	1.800 ₁₀ 1	2.721 ₁₀ -1	1.700 ₁₀ 1	2.938 ₁₀ -1
21	2.000 ₁₀ 1	2.449 ₁₀ -1	1.900 ₁₀ 1	2.631 ₁₀ -1	1.800 ₁₀ 1	2.831 ₁₀ -1
22	2.100 ₁₀ 1	2.381 ₁₀ -1	2.000 ₁₀ 1	2.549 ₁₀ -1	1.900 ₁₀ 1	2.734 ₁₀ -1
23	2.200 ₁₀ 1	2.318 ₁₀ -1	2.100 ₁₀ 1	2.474 ₁₀ -1	2.000 ₁₀ 1	2.645 ₁₀ -1
24	2.300 ₁₀ 1	2.259 ₁₀ -1	2.200 ₁₀ 1	2.405 ₁₀ -1	2.100 ₁₀ 1	2.564 ₁₀ -1
25	2.400 ₁₀ 1	2.205 ₁₀ -1	2.300 ₁₀ 1	2.341 ₁₀ -1	2.200 ₁₀ 1	2.489 ₁₀ -1
26	2.500 ₁₀ 1	2.154 ₁₀ -1	2.400 ₁₀ 1	2.282 ₁₀ -1	2.300 ₁₀ 1	2.421 ₁₀ -1
27	2.600 ₁₀ 1	2.107 ₁₀ -1	2.500 ₁₀ 1	2.227 ₁₀ -1	2.400 ₁₀ 1	2.357 ₁₀ -1
28	2.700 ₁₀ 1	2.062 ₁₀ -1	2.600 ₁₀ 1	2.176 ₁₀ -1	2.500 ₁₀ 1	2.298 ₁₀ -1
29	2.800 ₁₀ 1	2.020 ₁₀ -1	2.700 ₁₀ 1	2.128 ₁₀ -1	2.600 ₁₀ 1	2.243 ₁₀ -1
30	2.900 ₁₀ 1	1.981 ₁₀ -1	2.800 ₁₀ 1	2.082 ₁₀ -1	2.700 ₁₀ 1	2.191 ₁₀ -1
31	3.000 ₁₀ 1	1.944 ₁₀ -1	2.900 ₁₀ 1	2.040 ₁₀ -1	2.800 ₁₀ 1	2.143 ₁₀ -1
32	3.100 ₁₀ 1	1.908 ₁₀ -1	3.000 ₁₀ 1	2.000 ₁₀ -1	2.900 ₁₀ 1	2.098 ₁₀ -1
33	3.200 ₁₀ 1	1.875 ₁₀ -1	3.100 ₁₀ 1	1.962 ₁₀ -1	3.000 ₁₀ 1	2.055 ₁₀ -1
34	3.300 ₁₀ 1	1.843 ₁₀ -1	3.200 ₁₀ 1	1.926 ₁₀ -1	3.100 ₁₀ 1	2.015 ₁₀ -1
35	3.400 ₁₀ 1	1.813 ₁₀ -1	3.300 ₁₀ 1	1.892 ₁₀ -1	3.200 ₁₀ 1	1.976 ₁₀ -1
36	3.500 ₁₀ 1	1.784 ₁₀ -1	3.400 ₁₀ 1	1.860 ₁₀ -1	3.300 ₁₀ 1	1.940 ₁₀ -1
37	3.600 ₁₀ 1	1.757 ₁₀ -1	3.500 ₁₀ 1	1.829 ₁₀ -1	3.400 ₁₀ 1	1.906 ₁₀ -1
38	3.700 ₁₀ 1	1.731 ₁₀ -1	3.600 ₁₀ 1	1.800 ₁₀ -1	3.500 ₁₀ 1	1.874 ₁₀ -1
39	3.800 ₁₀ 1	1.705 ₁₀ -1	3.700 ₁₀ 1	1.772 ₁₀ -1	3.600 ₁₀ 1	1.843 ₁₀ -1
40	3.900 ₁₀ 1	1.681 ₁₀ -1	3.800 ₁₀ 1	1.746 ₁₀ -1	3.700 ₁₀ 1	1.813 ₁₀ -1
41	4.000 ₁₀ 1	1.658 ₁₀ -1	3.900 ₁₀ 1	1.720 ₁₀ -1	3.800 ₁₀ 1	1.785 ₁₀ -1
42	4.100 ₁₀ 1	1.636 ₁₀ -1	4.000 ₁₀ 1	1.696 ₁₀ -1	3.900 ₁₀ 1	1.758 ₁₀ -1
43	4.200 ₁₀ 1	1.615 ₁₀ -1	4.100 ₁₀ 1	1.672 ₁₀ -1	4.000 ₁₀ 1	1.732 ₁₀ -1
44	4.300 ₁₀ 1	1.594 ₁₀ -1	4.200 ₁₀ 1	1.650 ₁₀ -1	4.100 ₁₀ 1	1.707 ₁₀ -1
45	4.400 ₁₀ 1	1.575 ₁₀ -1	4.300 ₁₀ 1	1.628 ₁₀ -1	4.200 ₁₀ 1	1.684 ₁₀ -1
46	4.500 ₁₀ 1	1.556 ₁₀ -1	4.400 ₁₀ 1	1.607 ₁₀ -1	4.300 ₁₀ 1	1.661 ₁₀ -1
47	4.600 ₁₀ 1	1.537 ₁₀ -1	4.500 ₁₀ 1	1.587 ₁₀ -1	4.400 ₁₀ 1	1.639 ₁₀ -1
48	4.700 ₁₀ 1	1.519 ₁₀ -1	4.600 ₁₀ 1	1.568 ₁₀ -1	4.500 ₁₀ 1	1.618 ₁₀ -1
49	4.800 ₁₀ 1	1.502 ₁₀ -1	4.700 ₁₀ 1	1.549 ₁₀ -1	4.600 ₁₀ 1	1.597 ₁₀ -1
50	4.900 ₁₀ 1	1.486 ₁₀ -1	4.800 ₁₀ 1	1.531 ₁₀ -1	4.700 ₁₀ 1	1.578 ₁₀ -1
51	5.000 ₁₀ 1	1.470 ₁₀ -1	4.900 ₁₀ 1	1.514 ₁₀ -1	4.800 ₁₀ 1	1.559 ₁₀ -1

		n2 = 4		n2 = 5		n2 = 6	
n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	
1	1.286	9.493 ₁₀ -1	1.250	9.592 ₁₀ -1	1.222	9.664 ₁₀ -1	
2	1.621	9.009 ₁₀ -1	1.541	9.192 ₁₀ -1	1.478	9.329 ₁₀ -1	
3	2.011	8.536 ₁₀ -1	1.877	8.792 ₁₀ -1	1.773	8.989 ₁₀ -1	
4	2.461	8.070 ₁₀ -1	2.264	8.390 ₁₀ -1	2.110	8.642 ₁₀ -1	
5	2.975	7.610 ₁₀ -1	2.707	7.985 ₁₀ -1	2.496	8.286 ₁₀ -1	
6	3.555	7.159 ₁₀ -1	3.209	7.578 ₁₀ -1	2.933	7.924 ₁₀ -1	
7	4.199	6.720 ₁₀ -1	3.772	7.173 ₁₀ -1	3.425	7.556 ₁₀ -1	
8	4.906	6.299 ₁₀ -1	4.396	6.774 ₁₀ -1	3.974	7.186 ₁₀ -1	
9	5.671	5.899 ₁₀ -1	5.079	6.385 ₁₀ -1	4.581	6.818 ₁₀ -1	
10	6.487	5.524 ₁₀ -1	5.819	6.010 ₁₀ -1	5.246	6.455 ₁₀ -1	
11	7.347	5.175 ₁₀ -1	6.610	5.653 ₁₀ -1	5.965	6.102 ₁₀ -1	
12	8.242	4.854 ₁₀ -1	7.446	5.317 ₁₀ -1	6.734	5.762 ₁₀ -1	
13	9.166	4.561 ₁₀ -1	8.321	5.003 ₁₀ -1	7.549	5.438 ₁₀ -1	
14	1.011 ₁₀ 1	4.296 ₁₀ -1	9.226	4.713 ₁₀ -1	8.404	5.133 ₁₀ -1	
15	1.107 ₁₀ 1	4.056 ₁₀ -1	1.016 ₁₀ 1	4.447 ₁₀ -1	9.292	4.847 ₁₀ -1	
16	1.205 ₁₀ 1	3.841 ₁₀ -1	1.111 ₁₀ 1	4.204 ₁₀ -1	1.021 ₁₀ 1	4.582 ₁₀ -1	
17	1.303 ₁₀ 1	3.647 ₁₀ -1	1.207 ₁₀ 1	3.983 ₁₀ -1	1.115 ₁₀ 1	4.337 ₁₀ -1	
18	1.402 ₁₀ 1	3.474 ₁₀ -1	1.305 ₁₀ 1	3.783 ₁₀ -1	1.210 ₁₀ 1	4.113 ₁₀ -1	
19	1.501 ₁₀ 1	3.318 ₁₀ -1	1.403 ₁₀ 1	3.603 ₁₀ -1	1.307 ₁₀ 1	3.909 ₁₀ -1	
20	1.601 ₁₀ 1	3.178 ₁₀ -1	1.502 ₁₀ 1	3.439 ₁₀ -1	1.405 ₁₀ 1	3.723 ₁₀ -1	
21	1.700 ₁₀ 1	3.051 ₁₀ -1	1.601 ₁₀ 1	3.292 ₁₀ -1	1.503 ₁₀ 1	3.554 ₁₀ -1	
22	1.800 ₁₀ 1	2.936 ₁₀ -1	1.701 ₁₀ 1	3.158 ₁₀ -1	1.602 ₁₀ 1	3.400 ₁₀ -1	
23	1.900 ₁₀ 1	2.832 ₁₀ -1	1.801 ₁₀ 1	3.037 ₁₀ -1	1.701 ₁₀ 1	3.261 ₁₀ -1	
24	2.000 ₁₀ 1	2.737 ₁₀ -1	1.900 ₁₀ 1	2.927 ₁₀ -1	1.801 ₁₀ 1	3.134 ₁₀ -1	
25	2.100 ₁₀ 1	2.651 ₁₀ -1	2.000 ₁₀ 1	2.826 ₁₀ -1	1.901 ₁₀ 1	3.018 ₁₀ -1	
26	2.200 ₁₀ 1	2.571 ₁₀ -1	2.100 ₁₀ 1	2.734 ₁₀ -1	2.000 ₁₀ 1	2.912 ₁₀ -1	
27	2.300 ₁₀ 1	2.497 ₁₀ -1	2.200 ₁₀ 1	2.650 ₁₀ -1	2.100 ₁₀ 1	2.815 ₁₀ -1	
28	2.400 ₁₀ 1	2.429 ₁₀ -1	2.300 ₁₀ 1	2.572 ₁₀ -1	2.200 ₁₀ 1	2.726 ₁₀ -1	
29	2.500 ₁₀ 1	2.366 ₁₀ -1	2.400 ₁₀ 1	2.500 ₁₀ -1	2.300 ₁₀ 1	2.644 ₁₀ -1	
30	2.600 ₁₀ 1	2.308 ₁₀ -1	2.500 ₁₀ 1	2.433 ₁₀ -1	2.400 ₁₀ 1	2.568 ₁₀ -1	
31	2.700 ₁₀ 1	2.253 ₁₀ -1	2.600 ₁₀ 1	2.371 ₁₀ -1	2.500 ₁₀ 1	2.498 ₁₀ -1	
32	2.800 ₁₀ 1	2.202 ₁₀ -1	2.700 ₁₀ 1	2.313 ₁₀ -1	2.600 ₁₀ 1	2.432 ₁₀ -1	
33	2.900 ₁₀ 1	2.153 ₁₀ -1	2.800 ₁₀ 1	2.259 ₁₀ -1	2.700 ₁₀ 1	2.371 ₁₀ -1	
34	3.000 ₁₀ 1	2.108 ₁₀ -1	2.900 ₁₀ 1	2.208 ₁₀ -1	2.800 ₁₀ 1	2.314 ₁₀ -1	
35	3.100 ₁₀ 1	2.066 ₁₀ -1	3.000 ₁₀ 1	2.160 ₁₀ -1	2.900 ₁₀ 1	2.261 ₁₀ -1	
36	3.200 ₁₀ 1	2.025 ₁₀ -1	3.100 ₁₀ 1	2.115 ₁₀ -1	3.000 ₁₀ 1	2.211 ₁₀ -1	
37	3.300 ₁₀ 1	1.987 ₁₀ -1	3.200 ₁₀ 1	2.073 ₁₀ -1	3.100 ₁₀ 1	2.164 ₁₀ -1	
38	3.400 ₁₀ 1	1.951 ₁₀ -1	3.300 ₁₀ 1	2.033 ₁₀ -1	3.200 ₁₀ 1	2.119 ₁₀ -1	
39	3.500 ₁₀ 1	1.917 ₁₀ -1	3.400 ₁₀ 1	1.995 ₁₀ -1	3.300 ₁₀ 1	2.077 ₁₀ -1	
40	3.600 ₁₀ 1	1.884 ₁₀ -1	3.500 ₁₀ 1	1.959 ₁₀ -1	3.400 ₁₀ 1	2.038 ₁₀ -1	
41	3.700 ₁₀ 1	1.853 ₁₀ -1	3.600 ₁₀ 1	1.925 ₁₀ -1	3.500 ₁₀ 1	2.000 ₁₀ -1	
42	3.800 ₁₀ 1	1.823 ₁₀ -1	3.700 ₁₀ 1	1.892 ₁₀ -1	3.600 ₁₀ 1	1.964 ₁₀ -1	
43	3.900 ₁₀ 1	1.795 ₁₀ -1	3.800 ₁₀ 1	1.861 ₁₀ -1	3.700 ₁₀ 1	1.930 ₁₀ -1	
44	4.000 ₁₀ 1	1.768 ₁₀ -1	3.900 ₁₀ 1	1.831 ₁₀ -1	3.800 ₁₀ 1	1.898 ₁₀ -1	
45	4.100 ₁₀ 1	1.742 ₁₀ -1	4.000 ₁₀ 1	1.803 ₁₀ -1	3.900 ₁₀ 1	1.867 ₁₀ -1	
46	4.200 ₁₀ 1	1.717 ₁₀ -1	4.100 ₁₀ 1	1.776 ₁₀ -1	4.000 ₁₀ 1	1.837 ₁₀ -1	
47	4.300 ₁₀ 1	1.693 ₁₀ -1	4.200 ₁₀ 1	1.750 ₁₀ -1	4.100 ₁₀ 1	1.809 ₁₀ -1	
48	4.400 ₁₀ 1	1.670 ₁₀ -1	4.300 ₁₀ 1	1.725 ₁₀ -1	4.200 ₁₀ 1	1.782 ₁₀ -1	
49	4.500 ₁₀ 1	1.648 ₁₀ -1	4.400 ₁₀ 1	1.701 ₁₀ -1	4.300 ₁₀ 1	1.756 ₁₀ -1	
50	4.600 ₁₀ 1	1.627 ₁₀ -1	4.500 ₁₀ 1	1.678 ₁₀ -1	4.400 ₁₀ 1	1.731 ₁₀ -1	
51	4.700 ₁₀ 1	1.606 ₁₀ -1	4.600 ₁₀ 1	1.656 ₁₀ -1	4.500 ₁₀ 1	1.707 ₁₀ -1	

n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.200	9.718 ₁₀ -1	1.182	9.760 ₁₀ -1	1.167	9.794 ₁₀ -1
2	1.429	9.434 ₁₀ -1	1.388	9.517 ₁₀ -1	1.354	9.583 ₁₀ -1
3	1.690	9.143 ₁₀ -1	1.622	9.265 ₁₀ -1	1.566	9.363 ₁₀ -1
4	1.987	8.842 ₁₀ -1	1.887	9.003 ₁₀ -1	1.805	9.134 ₁₀ -1
5	2.326	8.531 ₁₀ -1	2.188	8.730 ₁₀ -1	2.074	8.893 ₁₀ -1
6	2.710	8.209 ₁₀ -1	2.528	8.445 ₁₀ -1	2.377	8.641 ₁₀ -1
7	3.142	7.878 ₁₀ -1	2.910	8.149 ₁₀ -1	2.717	8.376 ₁₀ -1
8	3.626	7.540 ₁₀ -1	3.338	7.842 ₁₀ -1	3.099	8.100 ₁₀ -1
9	4.165	7.198 ₁₀ -1	3.816	7.528 ₁₀ -1	3.524	7.814 ₁₀ -1
10	4.758	6.855 ₁₀ -1	4.345	7.208 ₁₀ -1	3.996	7.519 ₁₀ -1
11	5.406	6.514 ₁₀ -1	4.927	6.886 ₁₀ -1	4.517	7.218 ₁₀ -1
12	6.107	6.180 ₁₀ -1	5.560	6.564 ₁₀ -1	5.088	6.913 ₁₀ -1
13	6.857	5.855 ₁₀ -1	6.246	6.246 ₁₀ -1	5.710	6.608 ₁₀ -1
14	7.654	5.543 ₁₀ -1	6.980	5.936 ₁₀ -1	6.381	6.304 ₁₀ -1
15	8.490	5.246 ₁₀ -1	7.759	5.634 ₁₀ -1	7.100	6.006 ₁₀ -1
16	9.362	4.965 ₁₀ -1	8.579	5.345 ₁₀ -1	7.864	5.715 ₁₀ -1
17	1.026 ₁₀ 1	4.703 ₁₀ -1	9.435	5.071 ₁₀ -1	8.669	5.434 ₁₀ -1
18	1.119 ₁₀ 1	4.458 ₁₀ -1	1.032 ₁₀ 1	4.812 ₁₀ -1	9.511	5.166 ₁₀ -1
19	1.213 ₁₀ 1	4.233 ₁₀ -1	1.124 ₁₀ 1	4.569 ₁₀ -1	1.038 ₁₀ 1	4.910 ₁₀ -1
20	1.309 ₁₀ 1	4.025 ₁₀ -1	1.217 ₁₀ 1	4.343 ₁₀ -1	1.129 ₁₀ 1	4.670 ₁₀ -1
21	1.406 ₁₀ 1	3.835 ₁₀ -1	1.312 ₁₀ 1	4.133 ₁₀ -1	1.221 ₁₀ 1	4.444 ₁₀ -1
22	1.504 ₁₀ 1	3.662 ₁₀ -1	1.408 ₁₀ 1	3.940 ₁₀ -1	1.315 ₁₀ 1	4.234 ₁₀ -1
23	1.603 ₁₀ 1	3.503 ₁₀ -1	1.506 ₁₀ 1	3.763 ₁₀ -1	1.411 ₁₀ 1	4.039 ₁₀ -1
24	1.702 ₁₀ 1	3.358 ₁₀ -1	1.604 ₁₀ 1	3.600 ₁₀ -1	1.508 ₁₀ 1	3.859 ₁₀ -1
25	1.801 ₁₀ 1	3.226 ₁₀ -1	1.703 ₁₀ 1	3.451 ₁₀ -1	1.605 ₁₀ 1	3.693 ₁₀ -1
26	1.901 ₁₀ 1	3.105 ₁₀ -1	1.802 ₁₀ 1	3.314 ₁₀ -1	1.704 ₁₀ 1	3.540 ₁₀ -1
27	2.001 ₁₀ 1	2.995 ₁₀ -1	1.901 ₁₀ 1	3.189 ₁₀ -1	1.803 ₁₀ 1	3.399 ₁₀ -1
28	2.100 ₁₀ 1	2.893 ₁₀ -1	2.001 ₁₀ 1	3.074 ₁₀ -1	1.902 ₁₀ 1	3.270 ₁₀ -1
29	2.200 ₁₀ 1	2.800 ₁₀ -1	2.101 ₁₀ 1	2.969 ₁₀ -1	2.001 ₁₀ 1	3.151 ₁₀ -1
30	2.300 ₁₀ 1	2.714 ₁₀ -1	2.200 ₁₀ 1	2.872 ₁₀ -1	2.101 ₁₀ 1	3.042 ₁₀ -1
31	2.400 ₁₀ 1	2.634 ₁₀ -1	2.300 ₁₀ 1	2.782 ₁₀ -1	2.201 ₁₀ 1	2.941 ₁₀ -1
32	2.500 ₁₀ 1	2.561 ₁₀ -1	2.400 ₁₀ 1	2.699 ₁₀ -1	2.300 ₁₀ 1	2.848 ₁₀ -1
33	2.600 ₁₀ 1	2.492 ₁₀ -1	2.500 ₁₀ 1	2.622 ₁₀ -1	2.400 ₁₀ 1	2.762 ₁₀ -1
34	2.700 ₁₀ 1	2.428 ₁₀ -1	2.600 ₁₀ 1	2.551 ₁₀ -1	2.500 ₁₀ 1	2.682 ₁₀ -1
35	2.800 ₁₀ 1	2.369 ₁₀ -1	2.700 ₁₀ 1	2.484 ₁₀ -1	2.600 ₁₀ 1	2.608 ₁₀ -1
36	2.900 ₁₀ 1	2.313 ₁₀ -1	2.800 ₁₀ 1	2.422 ₁₀ -1	2.700 ₁₀ 1	2.539 ₁₀ -1
37	3.000 ₁₀ 1	2.261 ₁₀ -1	2.900 ₁₀ 1	2.364 ₁₀ -1	2.800 ₁₀ 1	2.474 ₁₀ -1
38	3.100 ₁₀ 1	2.211 ₁₀ -1	3.000 ₁₀ 1	2.309 ₁₀ -1	2.900 ₁₀ 1	2.414 ₁₀ -1
39	3.200 ₁₀ 1	2.165 ₁₀ -1	3.100 ₁₀ 1	2.258 ₁₀ -1	3.000 ₁₀ 1	2.357 ₁₀ -1
40	3.300 ₁₀ 1	2.121 ₁₀ -1	3.200 ₁₀ 1	2.210 ₁₀ -1	3.100 ₁₀ 1	2.304 ₁₀ -1
41	3.400 ₁₀ 1	2.080 ₁₀ -1	3.300 ₁₀ 1	2.164 ₁₀ -1	3.200 ₁₀ 1	2.253 ₁₀ -1
42	3.500 ₁₀ 1	2.040 ₁₀ -1	3.400 ₁₀ 1	2.121 ₁₀ -1	3.300 ₁₀ 1	2.206 ₁₀ -1
43	3.600 ₁₀ 1	2.003 ₁₀ -1	3.500 ₁₀ 1	2.080 ₁₀ -1	3.400 ₁₀ 1	2.161 ₁₀ -1
44	3.700 ₁₀ 1	1.968 ₁₀ -1	3.600 ₁₀ 1	2.041 ₁₀ -1	3.500 ₁₀ 1	2.119 ₁₀ -1
45	3.800 ₁₀ 1	1.934 ₁₀ -1	3.700 ₁₀ 1	2.004 ₁₀ -1	3.600 ₁₀ 1	2.079 ₁₀ -1
46	3.900 ₁₀ 1	1.902 ₁₀ -1	3.800 ₁₀ 1	1.969 ₁₀ -1	3.700 ₁₀ 1	2.040 ₁₀ -1
47	4.000 ₁₀ 1	1.871 ₁₀ -1	3.900 ₁₀ 1	1.936 ₁₀ -1	3.800 ₁₀ 1	2.004 ₁₀ -1
48	4.100 ₁₀ 1	1.841 ₁₀ -1	4.000 ₁₀ 1	1.904 ₁₀ -1	3.900 ₁₀ 1	1.970 ₁₀ -1
49	4.200 ₁₀ 1	1.813 ₁₀ -1	4.100 ₁₀ 1	1.873 ₁₀ -1	4.000 ₁₀ 1	1.936 ₁₀ -1
50	4.300 ₁₀ 1	1.786 ₁₀ -1	4.200 ₁₀ 1	1.844 ₁₀ -1	4.100 ₁₀ 1	1.905 ₁₀ -1
51	4.400 ₁₀ 1	1.760 ₁₀ -1	4.300 ₁₀ 1	1.816 ₁₀ -1	4.200 ₁₀ 1	1.875 ₁₀ -1

n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.154	9.821 ₁₀ -1	1.143	9.843 ₁₀ -1	1.133	9.861 ₁₀ -1
2	1.326	9.636 ₁₀ -1	1.302	9.680 ₁₀ -1	1.281	9.716 ₁₀ -1
3	1.519	9.444 ₁₀ -1	1.479	9.510 ₁₀ -1	1.445	9.565 ₁₀ -1
4	1.736	9.242 ₁₀ -1	1.677	9.331 ₁₀ -1	1.627	9.406 ₁₀ -1
5	1.979	9.029 ₁₀ -1	1.898	9.142 ₁₀ -1	1.829	9.238 ₁₀ -1
6	2.251	8.805 ₁₀ -1	2.145	8.942 ₁₀ -1	2.054	9.059 ₁₀ -1
7	2.556	8.568 ₁₀ -1	2.420	8.731 ₁₀ -1	2.304	8.869 ₁₀ -1
8	2.897	8.320 ₁₀ -1	2.727	8.508 ₁₀ -1	2.582	8.669 ₁₀ -1
9	3.278	8.060 ₁₀ -1	3.069	8.273 ₁₀ -1	2.890	8.456 ₁₀ -1
10	3.700	7.790 ₁₀ -1	3.448	8.027 ₁₀ -1	3.233	8.233 ₁₀ -1
11	4.167	7.512 ₁₀ -1	3.868	7.771 ₁₀ -1	3.612	7.998 ₁₀ -1
12	4.682	7.227 ₁₀ -1	4.331	7.506 ₁₀ -1	4.030	7.754 ₁₀ -1
13	5.244	6.937 ₁₀ -1	4.839	7.235 ₁₀ -1	4.489	7.502 ₁₀ -1
14	5.855	6.646 ₁₀ -1	5.394	6.959 ₁₀ -1	4.991	7.242 ₁₀ -1
15	6.514	6.355 ₁₀ -1	5.995	6.680 ₁₀ -1	5.539	6.978 ₁₀ -1
16	7.220	6.068 ₁₀ -1	6.643	6.401 ₁₀ -1	6.132	6.710 ₁₀ -1
17	7.970	5.787 ₁₀ -1	7.337	6.124 ₁₀ -1	6.770	6.442 ₁₀ -1
18	8.761	5.514 ₁₀ -1	8.074	5.852 ₁₀ -1	7.453	6.174 ₁₀ -1
19	9.589	5.251 ₁₀ -1	8.852	5.586 ₁₀ -1	8.178	5.910 ₁₀ -1
20	1.045 ₁₀ 1	5.000 ₁₀ -1	9.668	5.329 ₁₀ -1	8.944	5.652 ₁₀ -1
21	1.134 ₁₀ 1	4.762 ₁₀ -1	1.052 ₁₀ 1	5.083 ₁₀ -1	9.748	5.401 ₁₀ -1
22	1.225 ₁₀ 1	4.538 ₁₀ -1	1.139 ₁₀ 1	4.847 ₁₀ -1	1.058 ₁₀ 1	5.158 ₁₀ -1
23	1.319 ₁₀ 1	4.327 ₁₀ -1	1.230 ₁₀ 1	4.624 ₁₀ -1	1.145 ₁₀ 1	4.926 ₁₀ -1
24	1.413 ₁₀ 1	4.131 ₁₀ -1	1.322 ₁₀ 1	4.415 ₁₀ -1	1.235 ₁₀ 1	4.705 ₁₀ -1
25	1.510 ₁₀ 1	3.949 ₁₀ -1	1.416 ₁₀ 1	4.218 ₁₀ -1	1.326 ₁₀ 1	4.496 ₁₀ -1
26	1.607 ₁₀ 1	3.780 ₁₀ -1	1.512 ₁₀ 1	4.034 ₁₀ -1	1.419 ₁₀ 1	4.299 ₁₀ -1
27	1.705 ₁₀ 1	3.624 ₁₀ -1	1.609 ₁₀ 1	3.863 ₁₀ -1	1.514 ₁₀ 1	4.115 ₁₀ -1
28	1.803 ₁₀ 1	3.480 ₁₀ -1	1.706 ₁₀ 1	3.705 ₁₀ -1	1.610 ₁₀ 1	3.942 ₁₀ -1
29	1.902 ₁₀ 1	3.348 ₁₀ -1	1.804 ₁₀ 1	3.558 ₁₀ -1	1.708 ₁₀ 1	3.782 ₁₀ -1
30	2.002 ₁₀ 1	3.226 ₁₀ -1	1.903 ₁₀ 1	3.423 ₁₀ -1	1.805 ₁₀ 1	3.633 ₁₀ -1
31	2.101 ₁₀ 1	3.113 ₁₀ -1	2.002 ₁₀ 1	3.297 ₁₀ -1	1.904 ₁₀ 1	3.495 ₁₀ -1
32	2.201 ₁₀ 1	3.009 ₁₀ -1	2.101 ₁₀ 1	3.181 ₁₀ -1	2.003 ₁₀ 1	3.367 ₁₀ -1
33	2.300 ₁₀ 1	2.912 ₁₀ -1	2.201 ₁₀ 1	3.074 ₁₀ -1	2.102 ₁₀ 1	3.248 ₁₀ -1
34	2.400 ₁₀ 1	2.823 ₁₀ -1	2.301 ₁₀ 1	2.975 ₁₀ -1	2.201 ₁₀ 1	3.138 ₁₀ -1
35	2.500 ₁₀ 1	2.741 ₁₀ -1	2.400 ₁₀ 1	2.883 ₁₀ -1	2.301 ₁₀ 1	3.036 ₁₀ -1
36	2.600 ₁₀ 1	2.664 ₁₀ -1	2.500 ₁₀ 1	2.798 ₁₀ -1	2.401 ₁₀ 1	2.941 ₁₀ -1
37	2.700 ₁₀ 1	2.592 ₁₀ -1	2.600 ₁₀ 1	2.718 ₁₀ -1	2.500 ₁₀ 1	2.853 ₁₀ -1
38	2.800 ₁₀ 1	2.525 ₁₀ -1	2.700 ₁₀ 1	2.644 ₁₀ -1	2.600 ₁₀ 1	2.771 ₁₀ -1
39	2.900 ₁₀ 1	2.462 ₁₀ -1	2.800 ₁₀ 1	2.575 ₁₀ -1	2.700 ₁₀ 1	2.695 ₁₀ -1
40	3.000 ₁₀ 1	2.404 ₁₀ -1	2.900 ₁₀ 1	2.510 ₁₀ -1	2.800 ₁₀ 1	2.624 ₁₀ -1
41	3.100 ₁₀ 1	2.348 ₁₀ -1	3.000 ₁₀ 1	2.449 ₁₀ -1	2.900 ₁₀ 1	2.557 ₁₀ -1
42	3.200 ₁₀ 1	2.296 ₁₀ -1	3.100 ₁₀ 1	2.392 ₁₀ -1	3.000 ₁₀ 1	2.494 ₁₀ -1
43	3.300 ₁₀ 1	2.247 ₁₀ -1	3.200 ₁₀ 1	2.338 ₁₀ -1	3.100 ₁₀ 1	2.435 ₁₀ -1
44	3.400 ₁₀ 1	2.201 ₁₀ -1	3.300 ₁₀ 1	2.288 ₁₀ -1	3.200 ₁₀ 1	2.380 ₁₀ -1
45	3.500 ₁₀ 1	2.157 ₁₀ -1	3.400 ₁₀ 1	2.240 ₁₀ -1	3.300 ₁₀ 1	2.328 ₁₀ -1
46	3.600 ₁₀ 1	2.115 ₁₀ -1	3.500 ₁₀ 1	2.195 ₁₀ -1	3.400 ₁₀ 1	2.278 ₁₀ -1
47	3.700 ₁₀ 1	2.076 ₁₀ -1	3.600 ₁₀ 1	2.152 ₁₀ -1	3.500 ₁₀ 1	2.231 ₁₀ -1
48	3.800 ₁₀ 1	2.038 ₁₀ -1	3.700 ₁₀ 1	2.111 ₁₀ -1	3.600 ₁₀ 1	2.187 ₁₀ -1
49	3.900 ₁₀ 1	2.003 ₁₀ -1	3.800 ₁₀ 1	2.072 ₁₀ -1	3.700 ₁₀ 1	2.145 ₁₀ -1
50	4.000 ₁₀ 1	1.969 ₁₀ -1	3.900 ₁₀ 1	2.035 ₁₀ -1	3.800 ₁₀ 1	2.105 ₁₀ -1
51	4.100 ₁₀ 1	1.936 ₁₀ -1	4.000 ₁₀ 1	2.000 ₁₀ -1	3.900 ₁₀ 1	2.067 ₁₀ -1

n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.125	9.876 ₁₀ -1	1.118	9.889 ₁₀ -1	1.111	9.899 ₁₀ -1
2	1.263	9.747 ₁₀ -1	1.247	9.773 ₁₀ -1	1.233	9.795 ₁₀ -1
3	1.415	9.612 ₁₀ -1	1.389	9.651 ₁₀ -1	1.366	9.685 ₁₀ -1
4	1.583	9.469 ₁₀ -1	1.545	9.523 ₁₀ -1	1.511	9.570 ₁₀ -1
5	1.769	9.319 ₁₀ -1	1.717	9.388 ₁₀ -1	1.671	9.447 ₁₀ -1
6	1.975	9.158 ₁₀ -1	1.907	9.243 ₁₀ -1	1.847	9.317 ₁₀ -1
7	2.204	8.988 ₁₀ -1	2.117	9.090 ₁₀ -1	2.041	9.178 ₁₀ -1
8	2.457	8.807 ₁₀ -1	2.348	8.926 ₁₀ -1	2.254	9.030 ₁₀ -1
9	2.737	8.615 ₁₀ -1	2.604	8.752 ₁₀ -1	2.489	8.872 ₁₀ -1
10	3.047	8.412 ₁₀ -1	2.887	8.568 ₁₀ -1	2.747	8.705 ₁₀ -1
11	3.390	8.198 ₁₀ -1	3.199	8.373 ₁₀ -1	3.032	8.527 ₁₀ -1
12	3.769	7.974 ₁₀ -1	3.542	8.168 ₁₀ -1	3.345	8.339 ₁₀ -1
13	4.184	7.740 ₁₀ -1	3.920	7.952 ₁₀ -1	3.689	8.141 ₁₀ -1
14	4.640	7.498 ₁₀ -1	4.334	7.728 ₁₀ -1	4.066	7.933 ₁₀ -1
15	5.138	7.249 ₁₀ -1	4.787	7.495 ₁₀ -1	4.478	7.717 ₁₀ -1
16	5.679	6.995 ₁₀ -1	5.280	7.256 ₁₀ -1	4.929	7.493 ₁₀ -1
17	6.264	6.737 ₁₀ -1	5.816	7.011 ₁₀ -1	5.418	7.262 ₁₀ -1
18	6.894	6.478 ₁₀ -1	6.394	6.762 ₁₀ -1	5.948	7.025 ₁₀ -1
19	7.567	6.220 ₁₀ -1	7.015	6.512 ₁₀ -1	6.520	6.785 ₁₀ -1
20	8.281	5.963 ₁₀ -1	7.679	6.261 ₁₀ -1	7.134	6.542 ₁₀ -1
21	9.036	5.712 ₁₀ -1	8.383	6.012 ₁₀ -1	7.789	6.299 ₁₀ -1
22	9.828	5.466 ₁₀ -1	9.128	5.766 ₁₀ -1	8.484	6.056 ₁₀ -1
23	1.065 ₁₀ 1	5.228 ₁₀ -1	9.909	5.526 ₁₀ -1	9.219	5.817 ₁₀ -1
24	1.151 ₁₀ 1	4.999 ₁₀ -1	1.073 ₁₀ 1	5.293 ₁₀ -1	9.991	5.582 ₁₀ -1
25	1.239 ₁₀ 1	4.780 ₁₀ -1	1.157 ₁₀ 1	5.067 ₁₀ -1	1.080 ₁₀ 1	5.353 ₁₀ -1
26	1.330 ₁₀ 1	4.573 ₁₀ -1	1.245 ₁₀ 1	4.851 ₁₀ -1	1.163 ₁₀ 1	5.131 ₁₀ -1
27	1.423 ₁₀ 1	4.376 ₁₀ -1	1.334 ₁₀ 1	4.644 ₁₀ -1	1.250 ₁₀ 1	4.917 ₁₀ -1
28	1.517 ₁₀ 1	4.191 ₁₀ -1	1.426 ₁₀ 1	4.449 ₁₀ -1	1.339 ₁₀ 1	4.712 ₁₀ -1
29	1.613 ₁₀ 1	4.018 ₁₀ -1	1.520 ₁₀ 1	4.263 ₁₀ -1	1.430 ₁₀ 1	4.517 ₁₀ -1
30	1.709 ₁₀ 1	3.856 ₁₀ -1	1.615 ₁₀ 1	4.089 ₁₀ -1	1.523 ₁₀ 1	4.332 ₁₀ -1
31	1.807 ₁₀ 1	3.705 ₁₀ -1	1.711 ₁₀ 1	3.926 ₁₀ -1	1.617 ₁₀ 1	4.158 ₁₀ -1
32	1.905 ₁₀ 1	3.564 ₁₀ -1	1.808 ₁₀ 1	3.773 ₁₀ -1	1.713 ₁₀ 1	3.993 ₁₀ -1
33	2.003 ₁₀ 1	3.434 ₁₀ -1	1.906 ₁₀ 1	3.631 ₁₀ -1	1.810 ₁₀ 1	3.839 ₁₀ -1
34	2.102 ₁₀ 1	3.312 ₁₀ -1	2.004 ₁₀ 1	3.498 ₁₀ -1	1.907 ₁₀ 1	3.695 ₁₀ -1
35	2.202 ₁₀ 1	3.200 ₁₀ -1	2.103 ₁₀ 1	3.375 ₁₀ -1	2.005 ₁₀ 1	3.561 ₁₀ -1
36	2.301 ₁₀ 1	3.095 ₁₀ -1	2.202 ₁₀ 1	3.260 ₁₀ -1	2.104 ₁₀ 1	3.435 ₁₀ -1
37	2.401 ₁₀ 1	2.998 ₁₀ -1	2.301 ₁₀ 1	3.153 ₁₀ -1	2.203 ₁₀ 1	3.318 ₁₀ -1
38	2.501 ₁₀ 1	2.908 ₁₀ -1	2.401 ₁₀ 1	3.053 ₁₀ -1	2.302 ₁₀ 1	3.209 ₁₀ -1
39	2.600 ₁₀ 1	2.823 ₁₀ -1	2.501 ₁₀ 1	2.961 ₁₀ -1	2.401 ₁₀ 1	3.107 ₁₀ -1
40	2.700 ₁₀ 1	2.745 ₁₀ -1	2.600 ₁₀ 1	2.874 ₁₀ -1	2.501 ₁₀ 1	3.013 ₁₀ -1
41	2.800 ₁₀ 1	2.671 ₁₀ -1	2.700 ₁₀ 1	2.794 ₁₀ -1	2.601 ₁₀ 1	2.924 ₁₀ -1
42	2.900 ₁₀ 1	2.603 ₁₀ -1	2.800 ₁₀ 1	2.718 ₁₀ -1	2.700 ₁₀ 1	2.842 ₁₀ -1
43	3.000 ₁₀ 1	2.538 ₁₀ -1	2.900 ₁₀ 1	2.648 ₁₀ -1	2.800 ₁₀ 1	2.764 ₁₀ -1
44	3.100 ₁₀ 1	2.477 ₁₀ -1	3.000 ₁₀ 1	2.581 ₁₀ -1	2.900 ₁₀ 1	2.692 ₁₀ -1
45	3.200 ₁₀ 1	2.420 ₁₀ -1	3.100 ₁₀ 1	2.519 ₁₀ -1	3.000 ₁₀ 1	2.624 ₁₀ -1
46	3.300 ₁₀ 1	2.367 ₁₀ -1	3.200 ₁₀ 1	2.460 ₁₀ -1	3.100 ₁₀ 1	2.560 ₁₀ -1
47	3.400 ₁₀ 1	2.316 ₁₀ -1	3.300 ₁₀ 1	2.405 ₁₀ -1	3.200 ₁₀ 1	2.500 ₁₀ -1
48	3.500 ₁₀ 1	2.268 ₁₀ -1	3.400 ₁₀ 1	2.353 ₁₀ -1	3.300 ₁₀ 1	2.443 ₁₀ -1
49	3.600 ₁₀ 1	2.222 ₁₀ -1	3.500 ₁₀ 1	2.303 ₁₀ -1	3.400 ₁₀ 1	2.389 ₁₀ -1
50	3.700 ₁₀ 1	2.179 ₁₀ -1	3.600 ₁₀ 1	2.257 ₁₀ -1	3.500 ₁₀ 1	2.339 ₁₀ -1
51	3.800 ₁₀ 1	2.138 ₁₀ -1	3.700 ₁₀ 1	2.212 ₁₀ -1	3.600 ₁₀ 1	2.291 ₁₀ -1

n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.105	9.909 ₁₀ -1	1.100	9.917 ₁₀ -1	1.095	9.924 ₁₀ -1
2	1.220	9.814 ₁₀ -1	1.209	9.831 ₁₀ -1	1.198	9.845 ₁₀ -1
3	1.345	9.715 ₁₀ -1	1.327	9.740 ₁₀ -1	1.310	9.762 ₁₀ -1
4	1.482	9.610 ₁₀ -1	1.455	9.645 ₁₀ -1	1.431	9.675 ₁₀ -1
5	1.631	9.499 ₁₀ -1	1.595	9.543 ₁₀ -1	1.563	9.582 ₁₀ -1
6	1.795	9.380 ₁₀ -1	1.748	9.436 ₁₀ -1	1.706	9.484 ₁₀ -1
7	1.974	9.254 ₁₀ -1	1.915	9.321 ₁₀ -1	1.862	9.380 ₁₀ -1
8	2.171	9.120 ₁₀ -1	2.098	9.199 ₁₀ -1	2.032	9.268 ₁₀ -1
9	2.387	8.977 ₁₀ -1	2.298	9.068 ₁₀ -1	2.218	9.149 ₁₀ -1
10	2.625	8.824 ₁₀ -1	2.517	8.929 ₁₀ -1	2.421	9.022 ₁₀ -1
11	2.885	8.662 ₁₀ -1	2.757	8.781 ₁₀ -1	2.643	8.887 ₁₀ -1
12	3.172	8.490 ₁₀ -1	3.020	8.624 ₁₀ -1	2.885	8.742 ₁₀ -1
13	3.486	8.308 ₁₀ -1	3.308	8.457 ₁₀ -1	3.150	8.589 ₁₀ -1
14	3.830	8.117 ₁₀ -1	3.623	8.281 ₁₀ -1	3.440	8.427 ₁₀ -1
15	4.207	7.917 ₁₀ -1	3.968	8.096 ₁₀ -1	3.757	8.256 ₁₀ -1
16	4.618	7.708 ₁₀ -1	4.344	7.901 ₁₀ -1	4.102	8.076 ₁₀ -1
17	5.066	7.491 ₁₀ -1	4.754	7.699 ₁₀ -1	4.478	7.888 ₁₀ -1
18	5.552	7.267 ₁₀ -1	5.200	7.489 ₁₀ -1	4.886	7.692 ₁₀ -1
19	6.077	7.039 ₁₀ -1	5.682	7.273 ₁₀ -1	5.330	7.488 ₁₀ -1
20	6.643	6.806 ₁₀ -1	6.203	7.051 ₁₀ -1	5.809	7.278 ₁₀ -1
21	7.250	6.570 ₁₀ -1	6.764	6.825 ₁₀ -1	6.326	7.062 ₁₀ -1
22	7.897	6.334 ₁₀ -1	7.364	6.596 ₁₀ -1	6.881	6.843 ₁₀ -1
23	8.584	6.098 ₁₀ -1	8.004	6.366 ₁₀ -1	7.476	6.620 ₁₀ -1
24	9.310	5.864 ₁₀ -1	8.683	6.136 ₁₀ -1	8.109	6.396 ₁₀ -1
25	1.007 ₁₀ 1	5.634 ₁₀ -1	9.400	5.907 ₁₀ -1	8.781	6.171 ₁₀ -1
26	1.087 ₁₀ 1	5.409 ₁₀ -1	1.015 ₁₀ 1	5.682 ₁₀ -1	9.490	5.948 ₁₀ -1
27	1.170 ₁₀ 1	5.190 ₁₀ -1	1.094 ₁₀ 1	5.461 ₁₀ -1	1.024 ₁₀ 1	5.727 ₁₀ -1
28	1.255 ₁₀ 1	4.979 ₁₀ -1	1.176 ₁₀ 1	5.245 ₁₀ -1	1.101 ₁₀ 1	5.510 ₁₀ -1
29	1.343 ₁₀ 1	4.776 ₁₀ -1	1.261 ₁₀ 1	5.037 ₁₀ -1	1.183 ₁₀ 1	5.298 ₁₀ -1
30	1.434 ₁₀ 1	4.582 ₁₀ -1	1.348 ₁₀ 1	4.836 ₁₀ -1	1.266 ₁₀ 1	5.092 ₁₀ -1
31	1.526 ₁₀ 1	4.397 ₁₀ -1	1.438 ₁₀ 1	4.643 ₁₀ -1	1.353 ₁₀ 1	4.893 ₁₀ -1
32	1.620 ₁₀ 1	4.223 ₁₀ -1	1.529 ₁₀ 1	4.459 ₁₀ -1	1.442 ₁₀ 1	4.701 ₁₀ -1
33	1.715 ₁₀ 1	4.058 ₁₀ -1	1.623 ₁₀ 1	4.285 ₁₀ -1	1.533 ₁₀ 1	4.519 ₁₀ -1
34	1.811 ₁₀ 1	3.903 ₁₀ -1	1.717 ₁₀ 1	4.120 ₁₀ -1	1.625 ₁₀ 1	4.344 ₁₀ -1
35	1.908 ₁₀ 1	3.757 ₁₀ -1	1.813 ₁₀ 1	3.964 ₁₀ -1	1.720 ₁₀ 1	4.179 ₁₀ -1
36	2.006 ₁₀ 1	3.621 ₁₀ -1	1.910 ₁₀ 1	3.817 ₁₀ -1	1.815 ₁₀ 1	4.022 ₁₀ -1
37	2.104 ₁₀ 1	3.494 ₁₀ -1	2.007 ₁₀ 1	3.680 ₁₀ -1	1.911 ₁₀ 1	3.875 ₁₀ -1
38	2.203 ₁₀ 1	3.375 ₁₀ -1	2.105 ₁₀ 1	3.551 ₁₀ -1	2.008 ₁₀ 1	3.736 ₁₀ -1
39	2.302 ₁₀ 1	3.264 ₁₀ -1	2.204 ₁₀ 1	3.430 ₁₀ -1	2.106 ₁₀ 1	3.606 ₁₀ -1
40	2.402 ₁₀ 1	3.160 ₁₀ -1	2.303 ₁₀ 1	3.317 ₁₀ -1	2.205 ₁₀ 1	3.484 ₁₀ -1
41	2.501 ₁₀ 1	3.063 ₁₀ -1	2.402 ₁₀ 1	3.212 ₁₀ -1	2.303 ₁₀ 1	3.369 ₁₀ -1
42	2.601 ₁₀ 1	2.973 ₁₀ -1	2.501 ₁₀ 1	3.113 ₁₀ -1	2.402 ₁₀ 1	3.262 ₁₀ -1
43	2.701 ₁₀ 1	2.888 ₁₀ -1	2.601 ₁₀ 1	3.021 ₁₀ -1	2.502 ₁₀ 1	3.162 ₁₀ -1
44	2.800 ₁₀ 1	2.809 ₁₀ -1	2.701 ₁₀ 1	2.934 ₁₀ -1	2.601 ₁₀ 1	3.068 ₁₀ -1
45	2.900 ₁₀ 1	2.735 ₁₀ -1	2.801 ₁₀ 1	2.853 ₁₀ -1	2.701 ₁₀ 1	2.979 ₁₀ -1
46	3.000 ₁₀ 1	2.665 ₁₀ -1	2.900 ₁₀ 1	2.778 ₁₀ -1	2.801 ₁₀ 1	2.897 ₁₀ -1
47	3.100 ₁₀ 1	2.600 ₁₀ -1	3.000 ₁₀ 1	2.706 ₁₀ -1	2.900 ₁₀ 1	2.819 ₁₀ -1
48	3.200 ₁₀ 1	2.538 ₁₀ -1	3.100 ₁₀ 1	2.639 ₁₀ -1	3.000 ₁₀ 1	2.747 ₁₀ -1
49	3.300 ₁₀ 1	2.480 ₁₀ -1	3.200 ₁₀ 1	2.576 ₁₀ -1	3.100 ₁₀ 1	2.678 ₁₀ -1
50	3.400 ₁₀ 1	2.425 ₁₀ -1	3.300 ₁₀ 1	2.517 ₁₀ -1	3.200 ₁₀ 1	2.614 ₁₀ -1
51	3.500 ₁₀ 1	2.373 ₁₀ -1	3.100 ₁₀ 1	2.460 ₁₀ -1	3.300 ₁₀ 1	2.553 ₁₀ -1

n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.091	9.930 ₁₀ -1	1.087	9.936 ₁₀ -1	1.083	9.941 ₁₀ -1
2	1.189	9.858 ₁₀ -1	1.181	9.869 ₁₀ -1	1.173	9.879 ₁₀ -1
3	1.295	9.782 ₁₀ -1	1.281	9.799 ₁₀ -1	1.269	9.814 ₁₀ -1
4	1.410	9.702 ₁₀ -1	1.390	9.725 ₁₀ -1	1.372	9.746 ₁₀ -1
5	1.534	9.617 ₁₀ -1	1.507	9.647 ₁₀ -1	1.484	9.674 ₁₀ -1
6	1.668	9.527 ₁₀ -1	1.635	9.565 ₁₀ -1	1.604	9.598 ₁₀ -1
7	1.815	9.431 ₁₀ -1	1.772	9.477 ₁₀ -1	1.734	9.517 ₁₀ -1
8	1.974	9.329 ₁₀ -1	1.921	9.383 ₁₀ -1	1.874	9.431 ₁₀ -1
9	2.147	9.220 ₁₀ -1	2.083	9.283 ₁₀ -1	2.026	9.339 ₁₀ -1
10	2.336	9.104 ₁₀ -1	2.260	9.177 ₁₀ -1	2.191	9.241 ₁₀ -1
11	2.542	8.980 ₁₀ -1	2.451	9.063 ₁₀ -1	2.370	9.137 ₁₀ -1
12	2.766	8.848 ₁₀ -1	2.659	8.941 ₁₀ -1	2.564	9.025 ₁₀ -1
13	3.011	8.707 ₁₀ -1	2.886	8.812 ₁₀ -1	2.775	8.906 ₁₀ -1
14	3.278	8.558 ₁₀ -1	3.133	8.675 ₁₀ -1	3.004	8.780 ₁₀ -1
15	3.569	8.400 ₁₀ -1	3.402	8.529 ₁₀ -1	3.252	8.645 ₁₀ -1
16	3.886	8.234 ₁₀ -1	3.694	8.376 ₁₀ -1	3.523	8.503 ₁₀ -1
17	4.232	8.059 ₁₀ -1	4.013	8.213 ₁₀ -1	3.817	8.353 ₁₀ -1
18	4.607	7.876 ₁₀ -1	4.358	8.043 ₁₀ -1	4.136	8.194 ₁₀ -1
19	5.015	7.685 ₁₀ -1	4.734	7.864 ₁₀ -1	4.482	8.028 ₁₀ -1
20	5.456	7.487 ₁₀ -1	5.140	7.679 ₁₀ -1	4.857	7.854 ₁₀ -1
21	5.933	7.283 ₁₀ -1	5.580	7.486 ₁₀ -1	5.263	7.673 ₁₀ -1
22	6.446	7.073 ₁₀ -1	6.054	7.287 ₁₀ -1	5.700	7.485 ₁₀ -1
23	6.997	6.859 ₁₀ -1	6.563	7.083 ₁₀ -1	6.172	7.291 ₁₀ -1
24	7.586	6.642 ₁₀ -1	7.110	6.875 ₁₀ -1	6.678	7.092 ₁₀ -1
25	8.213	6.423 ₁₀ -1	7.694	6.663 ₁₀ -1	7.221	6.889 ₁₀ -1
26	8.877	6.204 ₁₀ -1	8.315	6.449 ₁₀ -1	7.800	6.682 ₁₀ -1
27	9.579	5.985 ₁₀ -1	8.973	6.235 ₁₀ -1	8.415	6.474 ₁₀ -1
28	1.032 ₁₀ 1	5.769 ₁₀ -1	9.668	6.021 ₁₀ -1	9.067	6.264 ₁₀ -1
29	1.109 ₁₀ 1	5.556 ₁₀ -1	1.040 ₁₀ 1	5.808 ₁₀ -1	9.755	6.054 ₁₀ -1
30	1.189 ₁₀ 1	5.347 ₁₀ -1	1.116 ₁₀ 1	5.599 ₁₀ -1	1.048 ₁₀ 1	5.846 ₁₀ -1
31	1.272 ₁₀ 1	5.144 ₁₀ -1	1.196 ₁₀ 1	5.393 ₁₀ -1	1.123 ₁₀ 1	5.640 ₁₀ -1
32	1.358 ₁₀ 1	4.947 ₁₀ -1	1.278 ₁₀ 1	5.193 ₁₀ -1	1.202 ₁₀ 1	5.437 ₁₀ -1
33	1.446 ₁₀ 1	4.757 ₁₀ -1	1.363 ₁₀ 1	4.998 ₁₀ -1	1.284 ₁₀ 1	5.239 ₁₀ -1
34	1.536 ₁₀ 1	4.575 ₁₀ -1	1.451 ₁₀ 1	4.810 ₁₀ -1	1.368 ₁₀ 1	5.046 ₁₀ -1
35	1.628 ₁₀ 1	4.401 ₁₀ -1	1.540 ₁₀ 1	4.629 ₁₀ -1	1.455 ₁₀ 1	4.860 ₁₀ -1
36	1.722 ₁₀ 1	4.236 ₁₀ -1	1.632 ₁₀ 1	4.455 ₁₀ -1	1.544 ₁₀ 1	4.680 ₁₀ -1
37	1.817 ₁₀ 1	4.079 ₁₀ -1	1.725 ₁₀ 1	4.290 ₁₀ -1	1.635 ₁₀ 1	4.508 ₁₀ -1
38	1.913 ₁₀ 1	3.931 ₁₀ -1	1.819 ₁₀ 1	4.133 ₁₀ -1	1.727 ₁₀ 1	4.343 ₁₀ -1
39	2.010 ₁₀ 1	3.791 ₁₀ -1	1.915 ₁₀ 1	3.984 ₁₀ -1	1.821 ₁₀ 1	4.185 ₁₀ -1
40	2.107 ₁₀ 1	3.659 ₁₀ -1	2.011 ₁₀ 1	3.844 ₁₀ -1	1.916 ₁₀ 1	4.036 ₁₀ -1
41	2.205 ₁₀ 1	3.536 ₁₀ -1	2.108 ₁₀ 1	3.711 ₁₀ -1	2.013 ₁₀ 1	3.895 ₁₀ -1
42	2.304 ₁₀ 1	3.420 ₁₀ -1	2.206 ₁₀ 1	3.586 ₁₀ -1	2.110 ₁₀ 1	3.762 ₁₀ -1
43	2.403 ₁₀ 1	3.311 ₁₀ -1	2.305 ₁₀ 1	3.469 ₁₀ -1	2.207 ₁₀ 1	3.636 ₁₀ -1
44	2.502 ₁₀ 1	3.209 ₁₀ -1	2.403 ₁₀ 1	3.359 ₁₀ -1	2.305 ₁₀ 1	3.517 ₁₀ -1
45	2.602 ₁₀ 1	3.113 ₁₀ -1	2.503 ₁₀ 1	3.255 ₁₀ -1	2.404 ₁₀ 1	3.405 ₁₀ -1
46	2.701 ₁₀ 1	3.024 ₁₀ -1	2.602 ₁₀ 1	3.158 ₁₀ -1	2.503 ₁₀ 1	3.301 ₁₀ -1
47	2.801 ₁₀ 1	2.939 ₁₀ -1	2.701 ₁₀ 1	3.067 ₁₀ -1	2.602 ₁₀ 1	3.202 ₁₀ -1
48	2.901 ₁₀ 1	2.860 ₁₀ -1	2.801 ₁₀ 1	2.981 ₁₀ -1	2.702 ₁₀ 1	3.109 ₁₀ -1
49	3.000 ₁₀ 1	2.786 ₁₀ -1	2.901 ₁₀ 1	2.901 ₁₀ -1	2.801 ₁₀ 1	3.022 ₁₀ -1
50	3.100 ₁₀ 1	2.716 ₁₀ -1	3.000 ₁₀ 1	2.825 ₁₀ -1	2.901 ₁₀ 1	2.940 ₁₀ -1
51	3.200 ₁₀ 1	2.650 ₁₀ -1	3.100 ₁₀ 1	2.754 ₁₀ -1	3.001 ₁₀ 1	2.863 ₁₀ -1

n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.080	9.945 ₁₀ -1	1.077	9.949 ₁₀ -1	1.074	9.952 ₁₀ -1
2	1.166	9.888 ₁₀ -1	1.159	9.896 ₁₀ -1	1.153	9.903 ₁₀ -1
3	1.257	9.828 ₁₀ -1	1.247	9.840 ₁₀ -1	1.237	9.851 ₁₀ -1
4	1.356	9.765 ₁₀ -1	1.341	9.782 ₁₀ -1	1.327	9.797 ₁₀ -1
5	1.462	9.698 ₁₀ -1	1.442	9.720 ₁₀ -1	1.424	9.739 ₁₀ -1
6	1.576	9.628 ₁₀ -1	1.550	9.655 ₁₀ -1	1.527	9.679 ₁₀ -1
7	1.699	9.553 ₁₀ -1	1.667	9.586 ₁₀ -1	1.638	9.615 ₁₀ -1
8	1.831	9.474 ₁₀ -1	1.792	9.512 ₁₀ -1	1.756	9.547 ₁₀ -1
9	1.974	9.389 ₁₀ -1	1.927	9.434 ₁₀ -1	1.884	9.474 ₁₀ -1
10	2.129	9.299 ₁₀ -1	2.073	9.351 ₁₀ -1	2.022	9.397 ₁₀ -1
11	2.296	9.203 ₁₀ -1	2.230	9.262 ₁₀ -1	2.170	9.315 ₁₀ -1
12	2.478	9.100 ₁₀ -1	2.400	9.167 ₁₀ -1	2.329	9.227 ₁₀ -1
13	2.674	8.990 ₁₀ -1	2.584	9.066 ₁₀ -1	2.502	9.134 ₁₀ -1
14	2.887	8.874 ₁₀ -1	2.783	8.958 ₁₀ -1	2.688	9.034 ₁₀ -1
15	3.119	8.750 ₁₀ -1	2.998	8.844 ₁₀ -1	2.889	8.928 ₁₀ -1
16	3.369	8.618 ₁₀ -1	3.231	8.722 ₁₀ -1	3.106	8.816 ₁₀ -1
17	3.641	8.479 ₁₀ -1	3.484	8.593 ₁₀ -1	3.342	8.696 ₁₀ -1
18	3.937	8.332 ₁₀ -1	3.757	8.456 ₁₀ -1	3.596	8.569 ₁₀ -1
19	4.256	8.177 ₁₀ -1	4.054	8.312 ₁₀ -1	3.871	8.436 ₁₀ -1
20	4.603	8.015 ₁₀ -1	4.374	8.161 ₁₀ -1	4.168	8.294 ₁₀ -1
21	4.978	7.845 ₁₀ -1	4.721	8.002 ₁₀ -1	4.490	8.146 ₁₀ -1
22	5.382	7.668 ₁₀ -1	5.095	7.836 ₁₀ -1	4.837	7.990 ₁₀ -1
23	5.818	7.485 ₁₀ -1	5.499	7.663 ₁₀ -1	5.211	7.828 ₁₀ -1
24	6.288	7.295 ₁₀ -1	5.934	7.484 ₁₀ -1	5.614	7.659 ₁₀ -1
25	6.791	7.101 ₁₀ -1	6.401	7.299 ₁₀ -1	6.047	7.484 ₁₀ -1
26	7.330	6.902 ₁₀ -1	6.902	7.109 ₁₀ -1	6.512	7.303 ₁₀ -1
27	7.904	6.701 ₁₀ -1	7.436	6.915 ₁₀ -1	7.010	7.117 ₁₀ -1
28	8.514	6.496 ₁₀ -1	8.007	6.718 ₁₀ -1	7.542	6.927 ₁₀ -1
29	9.160	6.291 ₁₀ -1	8.612	6.518 ₁₀ -1	8.108	6.734 ₁₀ -1
30	9.842	6.085 ₁₀ -1	9.253	6.316 ₁₀ -1	8.708	6.538 ₁₀ -1
31	1.056 ₁₀ 1	5.881 ₁₀ -1	9.928	6.115 ₁₀ -1	9.344	6.341 ₁₀ -1
32	1.131 ₁₀ 1	5.678 ₁₀ -1	1.064 ₁₀ 1	5.914 ₁₀ -1	1.001 ₁₀ 1	6.143 ₁₀ -1
33	1.209 ₁₀ 1	5.479 ₁₀ -1	1.138 ₁₀ 1	5.714 ₁₀ -1	1.072 ₁₀ 1	5.945 ₁₀ -1
34	1.290 ₁₀ 1	5.283 ₁₀ -1	1.215 ₁₀ 1	5.518 ₁₀ -1	1.145 ₁₀ 1	5.749 ₁₀ -1
35	1.373 ₁₀ 1	5.093 ₁₀ -1	1.296 ₁₀ 1	5.325 ₁₀ -1	1.222 ₁₀ 1	5.555 ₁₀ -1
36	1.460 ₁₀ 1	4.908 ₁₀ -1	1.379 ₁₀ 1	5.137 ₁₀ -1	1.302 ₁₀ 1	5.365 ₁₀ -1
37	1.548 ₁₀ 1	4.729 ₁₀ -1	1.464 ₁₀ 1	4.953 ₁₀ -1	1.384 ₁₀ 1	5.179 ₁₀ -1
38	1.638 ₁₀ 1	4.557 ₁₀ -1	1.552 ₁₀ 1	4.776 ₁₀ -1	1.469 ₁₀ 1	4.997 ₁₀ -1
39	1.730 ₁₀ 1	4.393 ₁₀ -1	1.642 ₁₀ 1	4.605 ₁₀ -1	1.556 ₁₀ 1	4.821 ₁₀ -1
40	1.824 ₁₀ 1	4.236 ₁₀ -1	1.733 ₁₀ 1	4.441 ₁₀ -1	1.645 ₁₀ 1	4.651 ₁₀ -1
41	1.918 ₁₀ 1	4.086 ₁₀ -1	1.826 ₁₀ 1	4.285 ₁₀ -1	1.736 ₁₀ 1	4.488 ₁₀ -1
42	2.014 ₁₀ 1	3.945 ₁₀ -1	1.921 ₁₀ 1	4.135 ₁₀ -1	1.829 ₁₀ 1	4.331 ₁₀ -1
43	2.111 ₁₀ 1	3.810 ₁₀ -1	2.016 ₁₀ 1	3.993 ₁₀ -1	1.923 ₁₀ 1	4.182 ₁₀ -1
44	2.208 ₁₀ 1	3.684 ₁₀ -1	2.112 ₁₀ 1	3.858 ₁₀ -1	2.018 ₁₀ 1	4.039 ₁₀ -1
45	2.306 ₁₀ 1	3.564 ₁₀ -1	2.209 ₁₀ 1	3.730 ₁₀ -1	2.114 ₁₀ 1	3.904 ₁₀ -1
46	2.405 ₁₀ 1	3.451 ₁₀ -1	2.307 ₁₀ 1	3.609 ₁₀ -1	2.211 ₁₀ 1	3.776 ₁₀ -1
47	2.503 ₁₀ 1	3.345 ₁₀ -1	2.405 ₁₀ 1	3.496 ₁₀ -1	2.308 ₁₀ 1	3.654 ₁₀ -1
48	2.603 ₁₀ 1	3.245 ₁₀ -1	2.504 ₁₀ 1	3.388 ₁₀ -1	2.406 ₁₀ 1	3.539 ₁₀ -1
49	2.702 ₁₀ 1	3.151 ₁₀ -1	2.603 ₁₀ 1	3.287 ₁₀ -1	2.505 ₁₀ 1	3.431 ₁₀ -1
50	2.801 ₁₀ 1	3.063 ₁₀ -1	2.702 ₁₀ 1	3.192 ₁₀ -1	2.603 ₁₀ 1	3.328 ₁₀ -1
51	2.901 ₁₀ 1	2.979 ₁₀ -1	2.802 ₁₀ 1	3.102 ₁₀ -1	2.703 ₁₀ 1	3.232 ₁₀ -1

n1	n2 = 25	n2 = 26	n2 = 27			
	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.071	9.955 ₁₀ -1	1.069	9.958 ₁₀ -1	1.067	9.961 ₁₀ -1
2	1.147	9.909 ₁₀ -1	1.142	9.915 ₁₀ -1	1.137	9.920 ₁₀ -1
3	1.228	9.861 ₁₀ -1	1.220	9.870 ₁₀ -1	1.212	9.878 ₁₀ -1
4	1.315	9.810 ₁₀ -1	1.303	9.822 ₁₀ -1	1.292	9.833 ₁₀ -1
5	1.407	9.757 ₁₀ -1	1.391	9.773 ₁₀ -1	1.377	9.787 ₁₀ -1
6	1.505	9.700 ₁₀ -1	1.485	9.720 ₁₀ -1	1.467	9.738 ₁₀ -1
7	1.611	9.641 ₁₀ -1	1.586	9.664 ₁₀ -1	1.563	9.686 ₁₀ -1
8	1.724	9.578 ₁₀ -1	1.694	9.606 ₁₀ -1	1.666	9.631 ₁₀ -1
9	1.845	9.510 ₁₀ -1	1.809	9.543 ₁₀ -1	1.775	9.573 ₁₀ -1
10	1.975	9.439 ₁₀ -1	1.932	9.477 ₁₀ -1	1.893	9.511 ₁₀ -1
11	2.115	9.363 ₁₀ -1	2.064	9.406 ₁₀ -1	2.018	9.445 ₁₀ -1
12	2.265	9.281 ₁₀ -1	2.207	9.331 ₁₀ -1	2.153	9.375 ₁₀ -1
13	2.427	9.195 ₁₀ -1	2.359	9.250 ₁₀ -1	2.297	9.300 ₁₀ -1
14	2.602	9.103 ₁₀ -1	2.524	9.165 ₁₀ -1	2.452	9.221 ₁₀ -1
15	2.790	9.005 ₁₀ -1	2.700	9.074 ₁₀ -1	2.618	9.137 ₁₀ -1
16	2.993	8.900 ₁₀ -1	2.891	8.977 ₁₀ -1	2.797	9.047 ₁₀ -1
17	3.213	8.790 ₁₀ -1	3.096	8.874 ₁₀ -1	2.990	8.951 ₁₀ -1
18	3.450	8.672 ₁₀ -1	3.317	8.765 ₁₀ -1	3.197	8.850 ₁₀ -1
19	3.706	8.547 ₁₀ -1	3.556	8.649 ₁₀ -1	3.420	8.742 ₁₀ -1
20	3.982	8.416 ₁₀ -1	3.814	8.527 ₁₀ -1	3.661	8.628 ₁₀ -1
21	4.281	8.277 ₁₀ -1	4.091	8.398 ₁₀ -1	3.920	8.508 ₁₀ -1
22	4.603	8.132 ₁₀ -1	4.391	8.262 ₁₀ -1	4.198	8.381 ₁₀ -1
23	4.950	7.980 ₁₀ -1	4.713	8.119 ₁₀ -1	4.499	8.247 ₁₀ -1
24	5.324	7.820 ₁₀ -1	5.061	7.969 ₁₀ -1	4.822	8.107 ₁₀ -1
25	5.726	7.655 ₁₀ -1	5.435	7.813 ₁₀ -1	5.170	7.960 ₁₀ -1
26	6.158	7.483 ₁₀ -1	5.836	7.651 ₁₀ -1	5.543	7.807 ₁₀ -1
27	6.621	7.306 ₁₀ -1	6.267	7.483 ₁₀ -1	5.944	7.648 ₁₀ -1
28	7.116	7.125 ₁₀ -1	6.728	7.310 ₁₀ -1	6.374	7.483 ₁₀ -1
29	7.645	6.939 ₁₀ -1	7.221	7.132 ₁₀ -1	6.833	7.313 ₁₀ -1
30	8.207	6.749 ₁₀ -1	7.746	6.949 ₁₀ -1	7.324	7.138 ₁₀ -1
31	8.803	6.557 ₁₀ -1	8.305	6.764 ₁₀ -1	7.846	6.959 ₁₀ -1
32	9.434	6.363 ₁₀ -1	8.897	6.575 ₁₀ -1	8.402	6.777 ₁₀ -1
33	1.010 ₁₀ 1	6.169 ₁₀ -1	9.523	6.385 ₁₀ -1	8.990	6.592 ₁₀ -1
34	1.080 ₁₀ 1	5.975 ₁₀ -1	1.018 ₁₀ 1	6.194 ₁₀ -1	9.611	6.406 ₁₀ -1
35	1.153 ₁₀ 1	5.782 ₁₀ -1	1.087 ₁₀ 1	6.003 ₁₀ -1	1.027 ₁₀ 1	6.218 ₁₀ -1
36	1.229 ₁₀ 1	5.591 ₁₀ -1	1.160 ₁₀ 1	5.813 ₁₀ -1	1.095 ₁₀ 1	6.030 ₁₀ -1
37	1.308 ₁₀ 1	5.403 ₁₀ -1	1.235 ₁₀ 1	5.625 ₁₀ -1	1.167 ₁₀ 1	5.843 ₁₀ -1
38	1.389 ₁₀ 1	5.219 ₁₀ -1	1.314 ₁₀ 1	5.439 ₁₀ -1	1.242 ₁₀ 1	5.657 ₁₀ -1
39	1.474 ₁₀ 1	5.039 ₁₀ -1	1.395 ₁₀ 1	5.257 ₁₀ -1	1.320 ₁₀ 1	5.474 ₁₀ -1
40	1.560 ₁₀ 1	4.865 ₁₀ -1	1.478 ₁₀ 1	5.079 ₁₀ -1	1.400 ₁₀ 1	5.294 ₁₀ -1
41	1.649 ₁₀ 1	4.696 ₁₀ -1	1.564 ₁₀ 1	4.906 ₁₀ -1	1.483 ₁₀ 1	5.118 ₁₀ -1
42	1.739 ₁₀ 1	4.533 ₁₀ -1	1.653 ₁₀ 1	4.738 ₁₀ -1	1.569 ₁₀ 1	4.946 ₁₀ -1
43	1.831 ₁₀ 1	4.377 ₁₀ -1	1.743 ₁₀ 1	4.576 ₁₀ -1	1.656 ₁₀ 1	4.779 ₁₀ -1
44	1.925 ₁₀ 1	4.227 ₁₀ -1	1.834 ₁₀ 1	4.421 ₁₀ -1	1.746 ₁₀ 1	4.618 ₁₀ -1
45	2.020 ₁₀ 1	4.085 ₁₀ -1	1.927 ₁₀ 1	4.271 ₁₀ -1	1.837 ₁₀ 1	4.463 ₁₀ -1
46	2.115 ₁₀ 1	3.949 ₁₀ -1	2.022 ₁₀ 1	4.128 ₁₀ -1	1.930 ₁₀ 1	4.314 ₁₀ -1
47	2.212 ₁₀ 1	3.820 ₁₀ -1	2.117 ₁₀ 1	3.992 ₁₀ -1	2.024 ₁₀ 1	4.171 ₁₀ -1
48	2.309 ₁₀ 1	3.697 ₁₀ -1	2.213 ₁₀ 1	3.863 ₁₀ -1	2.119 ₁₀ 1	4.034 ₁₀ -1
49	2.407 ₁₀ 1	3.582 ₁₀ -1	2.310 ₁₀ 1	3.740 ₁₀ -1	2.215 ₁₀ 1	3.905 ₁₀ -1
50	2.505 ₁₀ 1	3.472 ₁₀ -1	2.408 ₁₀ 1	3.623 ₁₀ -1	2.311 ₁₀ 1	3.781 ₁₀ -1
51	2.604 ₁₀ 1	3.369 ₁₀ -1	2.506 ₁₀ 1	3.513 ₁₀ -1	2.409 ₁₀ 1	3.664 ₁₀ -1

n1	n2 = 28		n2 = 29		n2 = 30	
	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.065	9.963 ₁₀ -1	1.062	9.965 ₁₀ -1	1.061	9.967 ₁₀ -1
2	1.133	9.925 ₁₀ -1	1.129	9.929 ₁₀ -1	1.125	9.933 ₁₀ -1
3	1.205	9.885 ₁₀ -1	1.198	9.892 ₁₀ -1	1.192	9.898 ₁₀ -1
4	1.282	9.844 ₁₀ -1	1.272	9.853 ₁₀ -1	1.264	9.861 ₁₀ -1
5	1.363	9.800 ₁₀ -1	1.351	9.812 ₁₀ -1	1.339	9.823 ₁₀ -1
6	1.450	9.754 ₁₀ -1	1.434	9.769 ₁₀ -1	1.419	9.782 ₁₀ -1
7	1.542	9.705 ₁₀ -1	1.522	9.723 ₁₀ -1	1.504	9.739 ₁₀ -1
8	1.640	9.654 ₁₀ -1	1.616	9.675 ₁₀ -1	1.594	9.694 ₁₀ -1
9	1.745	9.600 ₁₀ -1	1.716	9.624 ₁₀ -1	1.690	9.646 ₁₀ -1
10	1.856	9.542 ₁₀ -1	1.823	9.570 ₁₀ -1	1.792	9.596 ₁₀ -1
11	1.976	9.481 ₁₀ -1	1.936	9.513 ₁₀ -1	1.900	9.542 ₁₀ -1
12	2.103	9.415 ₁₀ -1	2.058	9.452 ₁₀ -1	2.016	9.486 ₁₀ -1
13	2.240	9.346 ₁₀ -1	2.187	9.387 ₁₀ -1	2.139	9.425 ₁₀ -1
14	2.386	9.272 ₁₀ -1	2.326	9.319 ₁₀ -1	2.270	9.361 ₁₀ -1
15	2.543	9.194 ₁₀ -1	2.474	9.245 ₁₀ -1	2.411	9.293 ₁₀ -1
16	2.712	9.110 ₁₀ -1	2.634	9.168 ₁₀ -1	2.561	9.220 ₁₀ -1
17	2.893	9.021 ₁₀ -1	2.804	9.085 ₁₀ -1	2.722	9.143 ₁₀ -1
18	3.087	8.927 ₁₀ -1	2.987	8.997 ₁₀ -1	2.895	9.061 ₁₀ -1
19	3.296	8.827 ₁₀ -1	3.183	8.904 ₁₀ -1	3.080	8.974 ₁₀ -1
20	3.521	8.720 ₁₀ -1	3.394	8.805 ₁₀ -1	3.278	8.882 ₁₀ -1
21	3.763	8.608 ₁₀ -1	3.621	8.700 ₁₀ -1	3.491	8.784 ₁₀ -1
22	4.024	8.489 ₁₀ -1	3.864	8.589 ₁₀ -1	3.719	8.681 ₁₀ -1
23	4.304	8.364 ₁₀ -1	4.126	8.472 ₁₀ -1	3.964	8.571 ₁₀ -1
24	4.605	8.233 ₁₀ -1	4.407	8.349 ₁₀ -1	4.227	8.456 ₁₀ -1
25	4.929	8.095 ₁₀ -1	4.709	8.220 ₁₀ -1	4.509	8.335 ₁₀ -1
26	5.277	7.951 ₁₀ -1	5.034	8.084 ₁₀ -1	4.812	8.207 ₁₀ -1
27	5.650	7.801 ₁₀ -1	5.382	7.943 ₁₀ -1	5.137	8.074 ₁₀ -1
28	6.050	7.645 ₁₀ -1	5.755	7.795 ₁₀ -1	5.485	7.935 ₁₀ -1
29	6.479	7.483 ₁₀ -1	6.155	7.642 ₁₀ -1	5.858	7.790 ₁₀ -1
30	6.937	7.316 ₁₀ -1	6.582	7.483 ₁₀ -1	6.257	7.639 ₁₀ -1
31	7.425	7.145 ₁₀ -1	7.038	7.319 ₁₀ -1	6.684	7.483 ₁₀ -1
32	7.945	6.969 ₁₀ -1	7.525	7.151 ₁₀ -1	7.139	7.322 ₁₀ -1
33	8.497	6.790 ₁₀ -1	8.042	6.978 ₁₀ -1	7.623	7.156 ₁₀ -1
34	9.081	6.609 ₁₀ -1	8.591	6.803 ₁₀ -1	8.138	6.987 ₁₀ -1
35	9.698	6.425 ₁₀ -1	9.172	6.624 ₁₀ -1	8.683	6.814 ₁₀ -1
36	1.035 ₁₀ 1	6.241 ₁₀ -1	9.785	6.444 ₁₀ -1	9.261	6.639 ₁₀ -1
37	1.103 ₁₀ 1	6.056 ₁₀ -1	1.043 ₁₀ 1	6.262 ₁₀ -1	9.870	6.462 ₁₀ -1
38	1.174 ₁₀ 1	5.871 ₁₀ -1	1.111 ₁₀ 1	6.080 ₁₀ -1	1.051 ₁₀ 1	6.283 ₁₀ -1
39	1.249 ₁₀ 1	5.688 ₁₀ -1	1.181 ₁₀ 1	5.898 ₁₀ -1	1.118 ₁₀ 1	6.103 ₁₀ -1
40	1.326 ₁₀ 1	5.507 ₁₀ -1	1.255 ₁₀ 1	5.718 ₁₀ -1	1.189 ₁₀ 1	5.924 ₁₀ -1
41	1.406 ₁₀ 1	5.329 ₁₀ -1	1.332 ₁₀ 1	5.539 ₁₀ -1	1.262 ₁₀ 1	5.746 ₁₀ -1
42	1.488 ₁₀ 1	5.155 ₁₀ -1	1.411 ₁₀ 1	5.363 ₁₀ -1	1.338 ₁₀ 1	5.569 ₁₀ -1
43	1.573 ₁₀ 1	4.984 ₁₀ -1	1.493 ₁₀ 1	5.190 ₁₀ -1	1.417 ₁₀ 1	5.395 ₁₀ -1
44	1.660 ₁₀ 1	4.819 ₁₀ -1	1.578 ₁₀ 1	5.021 ₁₀ -1	1.498 ₁₀ 1	5.224 ₁₀ -1
45	1.749 ₁₀ 1	4.658 ₁₀ -1	1.664 ₁₀ 1	4.857 ₁₀ -1	1.582 ₁₀ 1	5.057 ₁₀ -1
46	1.840 ₁₀ 1	4.504 ₁₀ -1	1.753 ₁₀ 1	4.697 ₁₀ -1	1.668 ₁₀ 1	4.894 ₁₀ -1
47	1.932 ₁₀ 1	4.355 ₁₀ -1	1.843 ₁₀ 1	4.543 ₁₀ -1	1.756 ₁₀ 1	4.735 ₁₀ -1
48	2.026 ₁₀ 1	4.212 ₁₀ -1	1.935 ₁₀ 1	4.395 ₁₀ -1	1.846 ₁₀ 1	4.582 ₁₀ -1
49	2.120 ₁₀ 1	4.076 ₁₀ -1	2.028 ₁₀ 1	4.252 ₁₀ -1	1.937 ₁₀ 1	4.434 ₁₀ -1
50	2.216 ₁₀ 1	3.945 ₁₀ -1	2.122 ₁₀ 1	4.116 ₁₀ -1	2.030 ₁₀ 1	4.291 ₁₀ -1
51	2.313 ₁₀ 1	3.821 ₁₀ -1	2.218 ₁₀ 1	3.985 ₁₀ -1	2.124 ₁₀ 1	4.154 ₁₀ -1

n1	n2 = 31	n2 = 32	n2 = 33			
	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.059	9.969 ₁₀ -1	1.057	9.971 ₁₀ -1	1.056	9.972 ₁₀ -1
2	1.121	9.937 ₁₀ -1	1.117	9.940 ₁₀ -1	1.114	9.944 ₁₀ -1
3	1.186	9.904 ₁₀ -1	1.181	9.909 ₁₀ -1	1.175	9.914 ₁₀ -1
4	1.255	9.869 ₁₀ -1	1.247	9.876 ₁₀ -1	1.240	9.883 ₁₀ -1
5	1.328	9.833 ₁₀ -1	1.318	9.842 ₁₀ -1	1.308	9.850 ₁₀ -1
6	1.405	9.794 ₁₀ -1	1.392	9.806 ₁₀ -1	1.380	9.816 ₁₀ -1
7	1.487	9.754 ₁₀ -1	1.471	9.768 ₁₀ -1	1.456	9.780 ₁₀ -1
8	1.573	9.712 ₁₀ -1	1.554	9.728 ₁₀ -1	1.536	9.743 ₁₀ -1
9	1.665	9.667 ₁₀ -1	1.642	9.686 ₁₀ -1	1.621	9.703 ₁₀ -1
10	1.763	9.620 ₁₀ -1	1.736	9.641 ₁₀ -1	1.711	9.661 ₁₀ -1
11	1.866	9.569 ₁₀ -1	1.835	9.594 ₁₀ -1	1.806	9.617 ₁₀ -1
12	1.977	9.516 ₁₀ -1	1.940	9.544 ₁₀ -1	1.906	9.570 ₁₀ -1
13	2.094	9.460 ₁₀ -1	2.052	9.491 ₁₀ -1	2.014	9.520 ₁₀ -1
14	2.219	9.400 ₁₀ -1	2.171	9.435 ₁₀ -1	2.127	9.468 ₁₀ -1
15	2.353	9.336 ₁₀ -1	2.298	9.376 ₁₀ -1	2.248	9.412 ₁₀ -1
16	2.495	9.268 ₁₀ -1	2.434	9.312 ₁₀ -1	2.377	9.353 ₁₀ -1
17	2.647	9.196 ₁₀ -1	2.578	9.245 ₁₀ -1	2.514	9.290 ₁₀ -1
18	2.810	9.120 ₁₀ -1	2.732	9.173 ₁₀ -1	2.660	9.223 ₁₀ -1
19	2.985	9.039 ₁₀ -1	2.897	9.098 ₁₀ -1	2.816	9.152 ₁₀ -1
20	3.171	8.952 ₁₀ -1	3.073	9.017 ₁₀ -1	2.983	9.077 ₁₀ -1
21	3.371	8.861 ₁₀ -1	3.261	8.932 ₁₀ -1	3.161	8.997 ₁₀ -1
22	3.585	8.765 ₁₀ -1	3.463	8.842 ₁₀ -1	3.351	8.912 ₁₀ -1
23	3.815	8.662 ₁₀ -1	3.679	8.746 ₁₀ -1	3.554	8.823 ₁₀ -1
24	4.062	8.554 ₁₀ -1	3.910	8.645 ₁₀ -1	3.771	8.728 ₁₀ -1
25	4.326	8.441 ₁₀ -1	4.158	8.538 ₁₀ -1	4.004	8.628 ₁₀ -1
26	4.609	8.321 ₁₀ -1	4.423	8.426 ₁₀ -1	4.253	8.523 ₁₀ -1
27	4.913	8.196 ₁₀ -1	4.708	8.308 ₁₀ -1	4.520	8.412 ₁₀ -1
28	5.238	8.064 ₁₀ -1	5.012	8.185 ₁₀ -1	4.805	8.296 ₁₀ -1
29	5.587	7.927 ₁₀ -1	5.338	8.055 ₁₀ -1	5.110	8.174 ₁₀ -1
30	5.960	7.784 ₁₀ -1	5.687	7.920 ₁₀ -1	5.436	8.046 ₁₀ -1
31	6.358	7.636 ₁₀ -1	6.060	7.780 ₁₀ -1	5.785	7.913 ₁₀ -1
32	6.784	7.483 ₁₀ -1	6.458	7.634 ₁₀ -1	6.158	7.775 ₁₀ -1
33	7.237	7.325 ₁₀ -1	6.882	7.483 ₁₀ -1	6.555	7.631 ₁₀ -1
34	7.720	7.162 ₁₀ -1	7.334	7.327 ₁₀ -1	6.979	7.483 ₁₀ -1
35	8.232	6.996 ₁₀ -1	7.815	7.167 ₁₀ -1	7.430	7.330 ₁₀ -1
36	8.775	6.826 ₁₀ -1	8.325	7.004 ₁₀ -1	7.909	7.172 ₁₀ -1
37	9.349	6.653 ₁₀ -1	8.865	6.837 ₁₀ -1	8.417	7.011 ₁₀ -1
38	9.954	6.479 ₁₀ -1	9.436	6.667 ₁₀ -1	8.955	6.847 ₁₀ -1
39	1.059 ₁₀ 1	6.302 ₁₀ -1	1.004 ₁₀ 1	6.495 ₁₀ -1	9.522	6.680 ₁₀ -1
40	1.126 ₁₀ 1	6.126 ₁₀ -1	1.067 ₁₀ 1	6.321 ₁₀ -1	1.012 ₁₀ 1	6.510 ₁₀ -1
41	1.196 ₁₀ 1	5.949 ₁₀ -1	1.133 ₁₀ 1	6.147 ₁₀ -1	1.075 ₁₀ 1	6.339 ₁₀ -1
42	1.268 ₁₀ 1	5.773 ₁₀ -1	1.203 ₁₀ 1	5.973 ₁₀ -1	1.141 ₁₀ 1	6.168 ₁₀ -1
43	1.344 ₁₀ 1	5.599 ₁₀ -1	1.275 ₁₀ 1	5.799 ₁₀ -1	1.210 ₁₀ 1	5.996 ₁₀ -1
44	1.422 ₁₀ 1	5.426 ₁₀ -1	1.350 ₁₀ 1	5.627 ₁₀ -1	1.282 ₁₀ 1	5.824 ₁₀ -1
45	1.503 ₁₀ 1	5.257 ₁₀ -1	1.428 ₁₀ 1	5.456 ₁₀ -1	1.356 ₁₀ 1	5.654 ₁₀ -1
46	1.587 ₁₀ 1	5.091 ₁₀ -1	1.508 ₁₀ 1	5.289 ₁₀ -1	1.434 ₁₀ 1	5.485 ₁₀ -1
47	1.672 ₁₀ 1	4.929 ₁₀ -1	1.591 ₁₀ 1	5.124 ₁₀ -1	1.514 ₁₀ 1	5.319 ₁₀ -1
48	1.760 ₁₀ 1	4.771 ₁₀ -1	1.676 ₁₀ 1	4.963 ₁₀ -1	1.596 ₁₀ 1	5.156 ₁₀ -1
49	1.849 ₁₀ 1	4.619 ₁₀ -1	1.763 ₁₀ 1	4.807 ₁₀ -1	1.680 ₁₀ 1	4.996 ₁₀ -1
50	1.940 ₁₀ 1	4.471 ₁₀ -1	1.852 ₁₀ 1	4.655 ₁₀ -1	1.767 ₁₀ 1	4.841 ₁₀ -1
51	2.032 ₁₀ 1	4.329 ₁₀ -1	1.943 ₁₀ 1	4.508 ₁₀ -1	1.855 ₁₀ 1	4.689 ₁₀ -1

n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.054	9.974 ₁₀ -1	1.053	9.975 ₁₀ -1	1.051	9.976 ₁₀ -1
2	1.111	9.946 ₁₀ -1	1.108	9.949 ₁₀ -1	1.105	9.952 ₁₀ -1
3	1.170	9.918 ₁₀ -1	1.166	9.922 ₁₀ -1	1.161	9.926 ₁₀ -1
4	1.233	9.889 ₁₀ -1	1.227	9.894 ₁₀ -1	1.220	9.899 ₁₀ -1
5	1.299	9.858 ₁₀ -1	1.290	9.865 ₁₀ -1	1.282	9.872 ₁₀ -1
6	1.369	9.826 ₁₀ -1	1.358	9.835 ₁₀ -1	1.348	9.843 ₁₀ -1
7	1.442	9.792 ₁₀ -1	1.429	9.803 ₁₀ -1	1.416	9.813 ₁₀ -1
8	1.519	9.756 ₁₀ -1	1.503	9.769 ₁₀ -1	1.488	9.781 ₁₀ -1
9	1.601	9.719 ₁₀ -1	1.582	9.734 ₁₀ -1	1.564	9.747 ₁₀ -1
10	1.687	9.680 ₁₀ -1	1.665	9.696 ₁₀ -1	1.644	9.712 ₁₀ -1
11	1.778	9.638 ₁₀ -1	1.753	9.657 ₁₀ -1	1.729	9.675 ₁₀ -1
12	1.875	9.594 ₁₀ -1	1.845	9.616 ₁₀ -1	1.818	9.636 ₁₀ -1
13	1.977	9.547 ₁₀ -1	1.944	9.572 ₁₀ -1	1.912	9.594 ₁₀ -1
14	2.086	9.498 ₁₀ -1	2.048	9.525 ₁₀ -1	2.012	9.551 ₁₀ -1
15	2.202	9.445 ₁₀ -1	2.158	9.476 ₁₀ -1	2.117	9.504 ₁₀ -1
16	2.324	9.390 ₁₀ -1	2.275	9.424 ₁₀ -1	2.229	9.455 ₁₀ -1
17	2.455	9.331 ₁₀ -1	2.399	9.368 ₁₀ -1	2.348	9.403 ₁₀ -1
18	2.594	9.268 ₁₀ -1	2.532	9.310 ₁₀ -1	2.474	9.348 ₁₀ -1
19	2.741	9.202 ₁₀ -1	2.672	9.247 ₁₀ -1	2.608	9.290 ₁₀ -1
20	2.899	9.131 ₁₀ -1	2.822	9.181 ₁₀ -1	2.750	9.228 ₁₀ -1
21	3.067	9.056 ₁₀ -1	2.981	9.111 ₁₀ -1	2.901	9.162 ₁₀ -1
22	3.247	8.977 ₁₀ -1	3.151	9.037 ₁₀ -1	3.062	9.092 ₁₀ -1
23	3.439	8.894 ₁₀ -1	3.332	8.959 ₁₀ -1	3.234	9.019 ₁₀ -1
24	3.643	8.805 ₁₀ -1	3.525	8.876 ₁₀ -1	3.416	8.941 ₁₀ -1
25	3.862	8.711 ₁₀ -1	3.732	8.788 ₁₀ -1	3.611	8.859 ₁₀ -1
26	4.096	8.613 ₁₀ -1	3.952	8.695 ₁₀ -1	3.819	8.772 ₁₀ -1
27	4.347	8.509 ₁₀ -1	4.188	8.597 ₁₀ -1	4.041	8.680 ₁₀ -1
28	4.614	8.399 ₁₀ -1	4.439	8.495 ₁₀ -1	4.278	8.583 ₁₀ -1
29	4.901	8.284 ₁₀ -1	4.708	8.387 ₁₀ -1	4.530	8.481 ₁₀ -1
30	5.206	8.164 ₁₀ -1	4.995	8.273 ₁₀ -1	4.800	8.374 ₁₀ -1
31	5.533	8.038 ₁₀ -1	5.301	8.154 ₁₀ -1	5.088	8.262 ₁₀ -1
32	5.882	7.907 ₁₀ -1	5.629	8.030 ₁₀ -1	5.395	8.145 ₁₀ -1
33	6.255	7.771 ₁₀ -1	5.978	7.901 ₁₀ -1	5.723	8.023 ₁₀ -1
34	6.652	7.629 ₁₀ -1	6.350	7.767 ₁₀ -1	6.072	7.895 ₁₀ -1
35	7.075	7.483 ₁₀ -1	6.747	7.627 ₁₀ -1	6.444	7.763 ₁₀ -1
36	7.524	7.332 ₁₀ -1	7.169	7.483 ₁₀ -1	6.841	7.625 ₁₀ -1
37	8.001	7.177 ₁₀ -1	7.617	7.335 ₁₀ -1	7.262	7.483 ₁₀ -1
38	8.507	7.019 ₁₀ -1	8.093	7.182 ₁₀ -1	7.709	7.337 ₁₀ -1
39	9.043	6.857 ₁₀ -1	8.597	7.026 ₁₀ -1	8.183	7.187 ₁₀ -1
40	9.608	6.692 ₁₀ -1	9.130	6.866 ₁₀ -1	8.685	7.033 ₁₀ -1
41	1.020 ₁₀ 1	6.525 ₁₀ -1	9.692	6.704 ₁₀ -1	9.216	6.875 ₁₀ -1
42	1.083 ₁₀ 1	6.357 ₁₀ -1	1.028 ₁₀ 1	6.540 ₁₀ -1	9.775	6.715 ₁₀ -1
43	1.148 ₁₀ 1	6.187 ₁₀ -1	1.091 ₁₀ 1	6.373 ₁₀ -1	1.036 ₁₀ 1	6.553 ₁₀ -1
44	1.217 ₁₀ 1	6.018 ₁₀ -1	1.156 ₁₀ 1	6.206 ₁₀ -1	1.098 ₁₀ 1	6.389 ₁₀ -1
45	1.288 ₁₀ 1	5.848 ₁₀ -1	1.224 ₁₀ 1	6.039 ₁₀ -1	1.163 ₁₀ 1	6.225 ₁₀ -1
46	1.362 ₁₀ 1	5.680 ₁₀ -1	1.295 ₁₀ 1	5.871 ₁₀ -1	1.231 ₁₀ 1	6.059 ₁₀ -1
47	1.439 ₁₀ 1	5.513 ₁₀ -1	1.368 ₁₀ 1	5.705 ₁₀ -1	1.301 ₁₀ 1	5.894 ₁₀ -1
48	1.519 ₁₀ 1	5.349 ₁₀ -1	1.445 ₁₀ 1	5.540 ₁₀ -1	1.375 ₁₀ 1	5.729 ₁₀ -1
49	1.600 ₁₀ 1	5.187 ₁₀ -1	1.524 ₁₀ 1	5.377 ₁₀ -1	1.450 ₁₀ 1	5.566 ₁₀ -1
50	1.685 ₁₀ 1	5.028 ₁₀ -1	1.605 ₁₀ 1	5.217 ₁₀ -1	1.529 ₁₀ 1	5.405 ₁₀ -1
51	1.771 ₁₀ 1	4.874 ₁₀ -1	1.689 ₁₀ 1	5.059 ₁₀ -1	1.610 ₁₀ 1	5.246 ₁₀ -1

n2 = 37		n2 = 38		n2 = 39	
n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]
1	1.050	9.977 ₁₀ -1	1.049	9.978 ₁₀ -1	1.048
2	1.102	9.954 ₁₀ -1	1.100	9.956 ₁₀ -1	1.097
3	1.157	9.930 ₁₀ -1	1.153	9.933 ₁₀ -1	1.149
4	1.214	9.904 ₁₀ -1	1.209	9.909 ₁₀ -1	1.204
5	1.275	9.878 ₁₀ -1	1.267	9.884 ₁₀ -1	1.261
6	1.338	9.850 ₁₀ -1	1.329	9.858 ₁₀ -1	1.320
7	1.404	9.822 ₁₀ -1	1.393	9.830 ₁₀ -1	1.382
8	1.474	9.792 ₁₀ -1	1.461	9.802 ₁₀ -1	1.448
9	1.547	9.760 ₁₀ -1	1.532	9.772 ₁₀ -1	1.517
10	1.625	9.727 ₁₀ -1	1.606	9.740 ₁₀ -1	1.589
11	1.706	9.691 ₁₀ -1	1.685	9.707 ₁₀ -1	1.665
12	1.792	9.654 ₁₀ -1	1.767	9.672 ₁₀ -1	1.744
13	1.882	9.615 ₁₀ -1	1.855	9.635 ₁₀ -1	1.829
14	1.978	9.574 ₁₀ -1	1.947	9.596 ₁₀ -1	1.917
15	2.079	9.530 ₁₀ -1	2.044	9.555 ₁₀ -1	2.010
16	2.187	9.484 ₁₀ -1	2.147	9.511 ₁₀ -1	2.109
17	2.300	9.435 ₁₀ -1	2.255	9.465 ₁₀ -1	2.213
18	2.420	9.383 ₁₀ -1	2.370	9.416 ₁₀ -1	2.323
19	2.548	9.329 ₁₀ -1	2.492	9.365 ₁₀ -1	2.440
20	2.683	9.270 ₁₀ -1	2.621	9.310 ₁₀ -1	2.563
21	2.827	9.209 ₁₀ -1	2.758	9.252 ₁₀ -1	2.693
22	2.980	9.143 ₁₀ -1	2.903	9.191 ₁₀ -1	2.832
23	3.142	9.074 ₁₀ -1	3.058	9.126 ₁₀ -1	2.979
24	3.315	9.001 ₁₀ -1	3.222	9.057 ₁₀ -1	3.135
25	3.500	8.924 ₁₀ -1	3.396	8.984 ₁₀ -1	3.300
26	3.696	8.842 ₁₀ -1	3.582	8.908 ₁₀ -1	3.476
27	3.905	8.756 ₁₀ -1	3.780	8.826 ₁₀ -1	3.664
28	4.129	8.665 ₁₀ -1	3.991	8.741 ₁₀ -1	3.863
29	4.367	8.569 ₁₀ -1	4.215	8.651 ₁₀ -1	4.075
30	4.620	8.469 ₁₀ -1	4.454	8.556 ₁₀ -1	4.301
31	4.891	8.363 ₁₀ -1	4.709	8.456 ₁₀ -1	4.541
32	5.180	8.252 ₁₀ -1	4.981	8.352 ₁₀ -1	4.797
33	5.488	8.136 ₁₀ -1	5.270	8.242 ₁₀ -1	5.070
34	5.816	8.016 ₁₀ -1	5.579	8.128 ₁₀ -1	5.360
35	6.165	7.890 ₁₀ -1	5.907	8.009 ₁₀ -1	5.669
36	6.537	7.759 ₁₀ -1	6.257	7.884 ₁₀ -1	5.998
37	6.933	7.623 ₁₀ -1	6.629	7.755 ₁₀ -1	6.348
38	7.353	7.483 ₁₀ -1	7.024	7.622 ₁₀ -1	6.719
39	7.800	7.339 ₁₀ -1	7.444	7.484 ₁₀ -1	7.114
40	8.272	7.191 ₁₀ -1	7.889	7.341 ₁₀ -1	7.533
41	8.773	7.039 ₁₀ -1	8.360	7.195 ₁₀ -1	7.977
42	9.301	6.884 ₁₀ -1	8.859	7.045 ₁₀ -1	8.447
43	9.858	6.726 ₁₀ -1	9.385	6.893 ₁₀ -1	8.944
44	1.044 ₁₀ 1	6.566 ₁₀ -1	9.940	6.737 ₁₀ -1	9.468
45	1.106 ₁₀ 1	6.405 ₁₀ -1	1.052 ₁₀ 1	6.579 ₁₀ -1	1.002 ₁₀ 1
46	1.170 ₁₀ 1	6.242 ₁₀ -1	1.113 ₁₀ 1	6.420 ₁₀ -1	1.060 ₁₀ 1
47	1.238 ₁₀ 1	6.079 ₁₀ -1	1.178 ₁₀ 1	6.259 ₁₀ -1	1.121 ₁₀ 1
48	1.308 ₁₀ 1	5.915 ₁₀ -1	1.244 ₁₀ 1	6.098 ₁₀ -1	1.185 ₁₀ 1
49	1.381 ₁₀ 1	5.753 ₁₀ -1	1.314 ₁₀ 1	5.936 ₁₀ -1	1.251 ₁₀ 1
50	1.456 ₁₀ 1	5.591 ₁₀ -1	1.387 ₁₀ 1	5.775 ₁₀ -1	1.321 ₁₀ 1
51	1.534 ₁₀ 1	5.431 ₁₀ -1	1.462 ₁₀ 1	5.615 ₁₀ -1	1.393 ₁₀ 1

n2 = 40		n2 = 41		n2 = 42	
n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]
1	1.047	9.980 ₁₀ -1	1.045	9.981 ₁₀ -1	1.044
2	1.095	9.960 ₁₀ -1	1.093	9.962 ₁₀ -1	1.091
3	1.146	9.939 ₁₀ -1	1.142	9.941 ₁₀ -1	1.139
4	1.199	9.917 ₁₀ -1	1.194	9.920 ₁₀ -1	1.189
5	1.254	9.894 ₁₀ -1	1.248	9.899 ₁₀ -1	1.242
6	1.312	9.870 ₁₀ -1	1.304	9.876 ₁₀ -1	1.297
7	1.372	9.846 ₁₀ -1	1.363	9.852 ₁₀ -1	1.354
8	1.436	9.820 ₁₀ -1	1.425	9.828 ₁₀ -1	1.414
9	1.503	9.793 ₁₀ -1	1.489	9.802 ₁₀ -1	1.476
10	1.572	9.764 ₁₀ -1	1.557	9.775 ₁₀ -1	1.542
11	1.646	9.734 ₁₀ -1	1.628	9.747 ₁₀ -1	1.611
12	1.723	9.703 ₁₀ -1	1.702	9.717 ₁₀ -1	1.683
13	1.804	9.670 ₁₀ -1	1.781	9.685 ₁₀ -1	1.759
14	1.889	9.635 ₁₀ -1	1.867	9.652 ₁₀ -1	1.838
15	1.979	9.598 ₁₀ -1	1.949	9.617 ₁₀ -1	1.921
16	2.074	9.559 ₁₀ -1	2.041	9.581 ₁₀ -1	2.009
17	2.174	9.518 ₁₀ -1	2.137	9.542 ₁₀ -1	2.102
18	2.279	9.475 ₁₀ -1	2.238	9.501 ₁₀ -1	2.199
19	2.391	9.429 ₁₀ -1	2.345	9.457 ₁₀ -1	2.302
20	2.509	9.380 ₁₀ -1	2.458	9.412 ₁₀ -1	2.410
21	2.633	9.329 ₁₀ -1	2.577	9.363 ₁₀ -1	2.524
22	2.765	9.275 ₁₀ -1	2.703	9.312 ₁₀ -1	2.645
23	2.905	9.217 ₁₀ -1	2.837	9.258 ₁₀ -1	2.772
24	3.054	9.156 ₁₀ -1	2.978	9.201 ₁₀ -1	2.907
25	3.211	9.092 ₁₀ -1	3.128	9.140 ₁₀ -1	3.050
26	3.378	9.024 ₁₀ -1	3.287	9.076 ₁₀ -1	3.201
27	3.556	8.952 ₁₀ -1	3.455	9.009 ₁₀ -1	3.362
28	3.744	8.877 ₁₀ -1	3.634	8.937 ₁₀ -1	3.531
29	3.945	8.797 ₁₀ -1	3.824	8.862 ₁₀ -1	3.712
30	4.158	8.713 ₁₀ -1	4.026	8.783 ₁₀ -1	3.903
31	4.385	8.624 ₁₀ -1	4.241	8.699 ₁₀ -1	4.106
32	4.627	8.531 ₁₀ -1	4.469	8.612 ₁₀ -1	4.322
33	4.884	8.433 ₁₀ -1	4.711	8.519 ₁₀ -1	4.552
34	5.157	8.331 ₁₀ -1	4.969	8.423 ₁₀ -1	4.795
35	5.448	8.224 ₁₀ -1	5.244	8.321 ₁₀ -1	5.054
36	5.758	8.112 ₁₀ -1	5.535	8.215 ₁₀ -1	5.329
37	6.087	7.996 ₁₀ -1	5.846	8.105 ₁₀ -1	5.622
38	6.437	7.874 ₁₀ -1	6.175	7.990 ₁₀ -1	5.932
39	6.809	7.749 ₁₀ -1	6.525	7.870 ₁₀ -1	6.262
40	7.203	7.618 ₁₀ -1	6.897	7.746 ₁₀ -1	6.612
41	7.622	7.484 ₁₀ -1	7.291	7.617 ₁₀ -1	6.984
42	8.065	7.345 ₁₀ -1	7.709	7.484 ₁₀ -1	7.378
43	8.533	7.203 ₁₀ -1	8.151	7.347 ₁₀ -1	7.795
44	9.029	7.057 ₁₀ -1	8.618	7.207 ₁₀ -1	8.236
45	9.551	6.909 ₁₀ -1	9.112	7.063 ₁₀ -1	8.702
46	1.010 ₁₀ 1	6.757 ₁₀ -1	9.632	6.916 ₁₀ -1	9.195
47	1.068 ₁₀ 1	6.603 ₁₀ -1	1.018 ₁₀ 1	6.767 ₁₀ -1	9.713
48	1.128 ₁₀ 1	6.448 ₁₀ -1	1.076 ₁₀ 1	6.615 ₁₀ -1	1.026 ₁₀ 1
49	1.192 ₁₀ 1	6.291 ₁₀ -1	1.136 ₁₀ 1	6.461 ₁₀ -1	1.083 ₁₀ 1
50	1.258 ₁₀ 1	6.133 ₁₀ -1	1.199 ₁₀ 1	6.306 ₁₀ -1	1.143 ₁₀ 1
51	1.327 ₁₀ 1	5.975 ₁₀ -1	1.265 ₁₀ 1	6.150 ₁₀ -1	1.206 ₁₀ 1

n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.043	9.983 ₁₀ -1	1.043	9.983 ₁₀ -1	1.042	9.984 ₁₀ -1
2	1.089	9.965 ₁₀ -1	1.087	9.966 ₁₀ -1	1.085	9.967 ₁₀ -1
3	1.136	9.946 ₁₀ -1	1.133	9.948 ₁₀ -1	1.130	9.950 ₁₀ -1
4	1.185	9.927 ₁₀ -1	1.181	9.930 ₁₀ -1	1.177	9.933 ₁₀ -1
5	1.236	9.907 ₁₀ -1	1.231	9.911 ₁₀ -1	1.226	9.915 ₁₀ -1
6	1.290	9.887 ₁₀ -1	1.283	9.891 ₁₀ -1	1.276	9.896 ₁₀ -1
7	1.345	9.865 ₁₀ -1	1.337	9.871 ₁₀ -1	1.329	9.876 ₁₀ -1
8	1.403	9.843 ₁₀ -1	1.394	9.849 ₁₀ -1	1.384	9.855 ₁₀ -1
9	1.464	9.819 ₁₀ -1	1.453	9.827 ₁₀ -1	1.442	9.834 ₁₀ -1
10	1.528	9.795 ₁₀ -1	1.515	9.803 ₁₀ -1	1.502	9.812 ₁₀ -1
11	1.595	9.769 ₁₀ -1	1.579	9.779 ₁₀ -1	1.565	9.788 ₁₀ -1
12	1.665	9.742 ₁₀ -1	1.647	9.753 ₁₀ -1	1.631	9.764 ₁₀ -1
13	1.738	9.713 ₁₀ -1	1.718	9.726 ₁₀ -1	1.699	9.738 ₁₀ -1
14	1.814	9.684 ₁₀ -1	1.792	9.698 ₁₀ -1	1.771	9.711 ₁₀ -1
15	1.895	9.652 ₁₀ -1	1.870	9.668 ₁₀ -1	1.846	9.683 ₁₀ -1
16	1.980	9.619 ₁₀ -1	1.952	9.637 ₁₀ -1	1.925	9.653 ₁₀ -1
17	2.069	9.584 ₁₀ -1	2.038	9.604 ₁₀ -1	2.008	9.621 ₁₀ -1
18	2.162	9.548 ₁₀ -1	2.128	9.569 ₁₀ -1	2.095	9.588 ₁₀ -1
19	2.261	9.509 ₁₀ -1	2.223	9.532 ₁₀ -1	2.187	9.554 ₁₀ -1
20	2.365	9.468 ₁₀ -1	2.323	9.493 ₁₀ -1	2.283	9.517 ₁₀ -1
21	2.474	9.425 ₁₀ -1	2.428	9.452 ₁₀ -1	2.384	9.478 ₁₀ -1
22	2.590	9.379 ₁₀ -1	2.539	9.409 ₁₀ -1	2.490	9.437 ₁₀ -1
23	2.712	9.331 ₁₀ -1	2.655	9.364 ₁₀ -1	2.602	9.394 ₁₀ -1
24	2.841	9.280 ₁₀ -1	2.779	9.316 ₁₀ -1	2.720	9.349 ₁₀ -1
25	2.977	9.226 ₁₀ -1	2.909	9.265 ₁₀ -1	2.845	9.301 ₁₀ -1
26	3.121	9.169 ₁₀ -1	3.047	9.211 ₁₀ -1	2.977	9.250 ₁₀ -1
27	3.274	9.109 ₁₀ -1	3.192	9.154 ₁₀ -1	3.116	9.197 ₁₀ -1
28	3.436	9.046 ₁₀ -1	3.346	9.095 ₁₀ -1	3.263	9.140 ₁₀ -1
29	3.607	8.979 ₁₀ -1	3.509	9.032 ₁₀ -1	3.418	9.081 ₁₀ -1
30	3.789	8.909 ₁₀ -1	3.682	8.965 ₁₀ -1	3.582	9.018 ₁₀ -1
31	3.981	8.834 ₁₀ -1	3.865	8.895 ₁₀ -1	3.756	8.952 ₁₀ -1
32	4.186	8.756 ₁₀ -1	4.059	8.821 ₁₀ -1	3.940	8.882 ₁₀ -1
33	4.403	8.674 ₁₀ -1	4.265	8.744 ₁₀ -1	4.136	8.809 ₁₀ -1
34	4.633	8.588 ₁₀ -1	4.483	8.662 ₁₀ -1	4.342	8.732 ₁₀ -1
35	4.878	8.497 ₁₀ -1	4.714	8.576 ₁₀ -1	4.562	8.651 ₁₀ -1
36	5.138	8.402 ₁₀ -1	4.960	8.487 ₁₀ -1	4.794	8.566 ₁₀ -1
37	5.414	8.303 ₁₀ -1	5.221	8.393 ₁₀ -1	5.041	8.476 ₁₀ -1
38	5.707	8.199 ₁₀ -1	5.498	8.294 ₁₀ -1	5.303	8.383 ₁₀ -1
39	6.018	8.091 ₁₀ -1	5.791	8.191 ₁₀ -1	5.580	8.286 ₁₀ -1
40	6.348	7.978 ₁₀ -1	6.103	8.084 ₁₀ -1	5.875	8.184 ₁₀ -1
41	6.699	7.861 ₁₀ -1	6.434	7.973 ₁₀ -1	6.187	8.078 ₁₀ -1
42	7.070	7.740 ₁₀ -1	6.784	7.857 ₁₀ -1	6.518	7.967 ₁₀ -1
43	7.464	7.614 ₁₀ -1	7.155	7.737 ₁₀ -1	6.868	7.853 ₁₀ -1
44	7.880	7.485 ₁₀ -1	7.548	7.613 ₁₀ -1	7.239	7.734 ₁₀ -1
45	8.320	7.351 ₁₀ -1	7.964	7.485 ₁₀ -1	7.632	7.612 ₁₀ -1
46	8.786	7.214 ₁₀ -1	8.404	7.353 ₁₀ -1	8.048	7.485 ₁₀ -1
47	9.276	7.074 ₁₀ -1	8.868	7.218 ₁₀ -1	8.486	7.355 ₁₀ -1
48	9.793	6.930 ₁₀ -1	9.357	7.079 ₁₀ -1	8.949	7.221 ₁₀ -1
49	1.034 ₁₀ 1	6.785 ₁₀ -1	9.872	6.937 ₁₀ -1	9.437	7.084 ₁₀ -1
50	1.091 ₁₀ 1	6.636 ₁₀ -1	1.041 ₁₀ 1	6.793 ₁₀ -1	9.950	6.944 ₁₀ -1
51	1.150 ₁₀ 1	6.486 ₁₀ -1	1.098 ₁₀ 1	6.647 ₁₀ -1	1.049 ₁₀ 1	6.801 ₁₀ -1

n2 = 46		n2 = 47		n2 = 48	
n1	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]
1	1.041	9.985 ₁₀ -1	1.040	9.985 ₁₀ -1	1.039
2	1.083	9.969 ₁₀ -1	1.082	9.970 ₁₀ -1	1.080
3	1.127	9.952 ₁₀ -1	1.125	9.954 ₁₀ -1	1.122
4	1.173	9.936 ₁₀ -1	1.169	9.938 ₁₀ -1	1.166
5	1.221	9.918 ₁₀ -1	1.216	9.921 ₁₀ -1	1.211
6	1.270	9.900 ₁₀ -1	1.264	9.904 ₁₀ -1	1.259
7	1.322	9.881 ₁₀ -1	1.315	9.886 ₁₀ -1	1.308
8	1.375	9.861 ₁₀ -1	1.367	9.867 ₁₀ -1	1.359
9	1.432	9.841 ₁₀ -1	1.422	9.847 ₁₀ -1	1.412
10	1.490	9.819 ₁₀ -1	1.478	9.827 ₁₀ -1	1.467
11	1.551	9.797 ₁₀ -1	1.538	9.805 ₁₀ -1	1.525
12	1.615	9.774 ₁₀ -1	1.600	9.783 ₁₀ -1	1.585
13	1.681	9.749 ₁₀ -1	1.664	9.760 ₁₀ -1	1.648
14	1.751	9.723 ₁₀ -1	1.732	9.735 ₁₀ -1	1.714
15	1.824	9.696 ₁₀ -1	1.803	9.709 ₁₀ -1	1.782
16	1.900	9.668 ₁₀ -1	1.877	9.682 ₁₀ -1	1.854
17	1.980	9.638 ₁₀ -1	1.954	9.654 ₁₀ -1	1.929
18	2.064	9.607 ₁₀ -1	2.035	9.624 ₁₀ -1	2.007
19	2.153	9.574 ₁₀ -1	2.120	9.593 ₁₀ -1	2.090
20	2.245	9.539 ₁₀ -1	2.210	9.560 ₁₀ -1	2.176
21	2.342	9.502 ₁₀ -1	2.303	9.525 ₁₀ -1	2.266
22	2.445	9.464 ₁₀ -1	2.402	9.488 ₁₀ -1	2.361
23	2.552	9.423 ₁₀ -1	2.505	9.449 ₁₀ -1	2.460
24	2.666	9.380 ₁₀ -1	2.614	9.409 ₁₀ -1	2.565
25	2.785	9.334 ₁₀ -1	2.728	9.366 ₁₀ -1	2.675
26	2.911	9.286 ₁₀ -1	2.849	9.320 ₁₀ -1	2.791
27	3.044	9.236 ₁₀ -1	2.976	9.272 ₁₀ -1	2.913
28	3.184	9.182 ₁₀ -1	3.110	9.222 ₁₀ -1	3.041
29	3.332	9.126 ₁₀ -1	3.252	9.169 ₁₀ -1	3.177
30	3.489	9.067 ₁₀ -1	3.401	9.113 ₁₀ -1	3.319
31	3.654	9.004 ₁₀ -1	3.559	9.054 ₁₀ -1	3.470
32	3.830	8.939 ₁₀ -1	3.726	8.991 ₁₀ -1	3.629
33	4.015	8.869 ₁₀ -1	3.903	8.926 ₁₀ -1	3.797
34	4.212	8.796 ₁₀ -1	4.089	8.857 ₁₀ -1	3.975
35	4.420	8.720 ₁₀ -1	4.287	8.785 ₁₀ -1	4.163
36	4.640	8.639 ₁₀ -1	4.496	8.708 ₁₀ -1	4.361
37	4.874	8.555 ₁₀ -1	4.718	8.629 ₁₀ -1	4.572
38	5.121	8.467 ₁₀ -1	4.952	8.545 ₁₀ -1	4.794
39	5.384	8.374 ₁₀ -1	5.201	8.457 ₁₀ -1	5.030
40	5.662	8.278 ₁₀ -1	5.464	8.365 ₁₀ -1	5.279
41	5.957	8.177 ₁₀ -1	5.743	8.270 ₁₀ -1	5.544
42	6.270	8.072 ₁₀ -1	6.039	8.170 ₁₀ -1	5.823
43	6.601	7.962 ₁₀ -1	6.352	8.066 ₁₀ -1	6.120
44	6.952	7.849 ₁₀ -1	6.683	7.958 ₁₀ -1	6.433
45	7.323	7.732 ₁₀ -1	7.034	7.845 ₁₀ -1	6.765
46	7.715	7.610 ₁₀ -1	7.405	7.729 ₁₀ -1	7.116
47	8.130	7.485 ₁₀ -1	7.797	7.609 ₁₀ -1	7.487
48	8.568	7.356 ₁₀ -1	8.212	7.485 ₁₀ -1	7.879
49	9.030	7.224 ₁₀ -1	8.649	7.358 ₁₀ -1	8.292
50	9.516	7.089 ₁₀ -1	9.110	7.227 ₁₀ -1	8.729
51	1.003 ₁₀ 1	6.950 ₁₀ -1	9.595	7.093 ₁₀ -1	9.189

n1	n2 = 49			n2 = 50		
	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]	E[a]	D[a]/E[a]
1	1.038	9.986 ₁₀ -1	1.038	9.987 ₁₀ -1	1.037	9.987 ₁₀ -1
2	1.078	9.972 ₁₀ -1	1.077	9.973 ₁₀ -1	1.075	9.974 ₁₀ -1
3	1.120	9.958 ₁₀ -1	1.117	9.959 ₁₀ -1	1.115	9.961 ₁₀ -1
4	1.163	9.943 ₁₀ -1	1.159	9.945 ₁₀ -1	1.156	9.947 ₁₀ -1
5	1.207	9.927 ₁₀ -1	1.203	9.930 ₁₀ -1	1.199	9.932 ₁₀ -1
6	1.253	9.911 ₁₀ -1	1.248	9.914 ₁₀ -1	1.243	9.917 ₁₀ -1
7	1.301	9.894 ₁₀ -1	1.295	9.898 ₁₀ -1	1.289	9.902 ₁₀ -1
8	1.351	9.877 ₁₀ -1	1.344	9.882 ₁₀ -1	1.336	9.886 ₁₀ -1
9	1.403	9.859 ₁₀ -1	1.394	9.864 ₁₀ -1	1.386	9.869 ₁₀ -1
10	1.457	9.840 ₁₀ -1	1.447	9.846 ₁₀ -1	1.437	9.852 ₁₀ -1
11	1.513	9.821 ₁₀ -1	1.502	9.828 ₁₀ -1	1.491	9.834 ₁₀ -1
12	1.572	9.800 ₁₀ -1	1.559	9.808 ₁₀ -1	1.546	9.815 ₁₀ -1
13	1.633	9.779 ₁₀ -1	1.618	9.787 ₁₀ -1	1.604	9.796 ₁₀ -1
14	1.697	9.756 ₁₀ -1	1.680	9.766 ₁₀ -1	1.664	9.775 ₁₀ -1
15	1.763	9.733 ₁₀ -1	1.745	9.744 ₁₀ -1	1.727	9.754 ₁₀ -1
16	1.833	9.708 ₁₀ -1	1.812	9.720 ₁₀ -1	1.793	9.731 ₁₀ -1
17	1.905	9.683 ₁₀ -1	1.883	9.696 ₁₀ -1	1.861	9.708 ₁₀ -1
18	1.981	9.655 ₁₀ -1	1.956	9.670 ₁₀ -1	1.932	9.683 ₁₀ -1
19	2.061	9.627 ₁₀ -1	2.033	9.643 ₁₀ -1	2.007	9.657 ₁₀ -1
20	2.144	9.597 ₁₀ -1	2.113	9.614 ₁₀ -1	2.085	9.630 ₁₀ -1
21	2.231	9.566 ₁₀ -1	2.198	9.584 ₁₀ -1	2.166	9.602 ₁₀ -1
22	2.322	9.533 ₁₀ -1	2.286	9.553 ₁₀ -1	2.251	9.572 ₁₀ -1
23	2.418	9.498 ₁₀ -1	2.378	9.520 ₁₀ -1	2.341	9.540 ₁₀ -1
24	2.519	9.461 ₁₀ -1	2.475	9.485 ₁₀ -1	2.434	9.507 ₁₀ -1
25	2.625	9.422 ₁₀ -1	2.577	9.448 ₁₀ -1	2.532	9.472 ₁₀ -1
26	2.736	9.382 ₁₀ -1	2.684	9.409 ₁₀ -1	2.635	9.435 ₁₀ -1
27	2.853	9.339 ₁₀ -1	2.796	9.369 ₁₀ -1	2.743	9.397 ₁₀ -1
28	2.976	9.293 ₁₀ -1	2.914	9.326 ₁₀ -1	2.856	9.356 ₁₀ -1
29	3.106	9.246 ₁₀ -1	3.039	9.281 ₁₀ -1	2.976	9.313 ₁₀ -1
30	3.242	9.196 ₁₀ -1	3.170	9.233 ₁₀ -1	3.101	9.268 ₁₀ -1
31	3.386	9.143 ₁₀ -1	3.307	9.183 ₁₀ -1	3.233	9.221 ₁₀ -1
32	3.538	9.087 ₁₀ -1	3.452	9.130 ₁₀ -1	3.372	9.171 ₁₀ -1
33	3.698	9.028 ₁₀ -1	3.606	9.075 ₁₀ -1	3.518	9.118 ₁₀ -1
34	3.868	8.967 ₁₀ -1	3.767	9.016 ₁₀ -1	3.673	9.063 ₁₀ -1
35	4.046	8.902 ₁₀ -1	3.938	8.955 ₁₀ -1	3.835	9.004 ₁₀ -1
36	4.236	8.834 ₁₀ -1	4.118	8.890 ₁₀ -1	4.007	8.943 ₁₀ -1
37	4.435	8.762 ₁₀ -1	4.308	8.822 ₁₀ -1	4.188	8.879 ₁₀ -1
38	4.647	8.687 ₁₀ -1	4.509	8.751 ₁₀ -1	4.379	8.811 ₁₀ -1
39	4.870	8.608 ₁₀ -1	4.721	8.676 ₁₀ -1	4.581	8.741 ₁₀ -1
40	5.107	8.525 ₁₀ -1	4.946	8.598 ₁₀ -1	4.795	8.666 ₁₀ -1
41	5.357	8.439 ₁₀ -1	5.183	8.516 ₁₀ -1	5.020	8.589 ₁₀ -1
42	5.622	8.349 ₁₀ -1	5.435	8.430 ₁₀ -1	5.259	8.507 ₁₀ -1
43	5.903	8.255 ₁₀ -1	5.700	8.341 ₁₀ -1	5.511	8.422 ₁₀ -1
44	6.199	8.157 ₁₀ -1	5.981	8.247 ₁₀ -1	5.777	8.333 ₁₀ -1
45	6.513	8.054 ₁₀ -1	6.279	8.150 ₁₀ -1	6.059	8.241 ₁₀ -1
46	6.845	7.948 ₁₀ -1	6.593	8.149 ₁₀ -1	6.357	8.144 ₁₀ -1
47	7.196	7.838 ₁₀ -1	6.925	7.944 ₁₀ -1	6.672	8.044 ₁₀ -1
48	7.567	7.725 ₁₀ -1	7.276	7.835 ₁₀ -1	7.004	7.940 ₁₀ -1
49	7.959	7.607 ₁₀ -1	7.647	7.722 ₁₀ -1	7.355	7.832 ₁₀ -1
50	8.372	7.486 ₁₀ -1	8.039	7.606 ₁₀ -1	7.726	7.720 ₁₀ -1
51	8.808	7.361 ₁₀ -1	8.451	7.486 ₁₀ -1	8.117	7.605 ₁₀ -1

References

- 1) E. Rutherford and H. Geiger, with a note by H. Bateman, The Probability Variations in the Distribution of α -Particles.
Phil. Mag. 20, 698 (July-Dec. 1910).
- 2) H. Cramer, Mathematical Methods of Statistics (Princeton University Press, Seventh Printing, 1957).
- 3) M. Annis, W. Cheston and H. Primakoff, On Statistical Estimation in Physics.
Rev. Mod. Phys. 25, 818 (1957).
- 4) A. Ruark and L. Devol, The General Theory of Fluctuations in Radioactive Disintegration.
Phys. Rev. 49, 355 (1936).
- 5) L. J. Rainwater and C. S. Wu, Applications of Probability Theory to Nuclear Particle Detection.
Nucleonics 1, 60 (Oct. 1947).
- 6) Standard Mathematical Tables (Chemical Rubber Publishing Company, Cleveland, Ohio, 1954).